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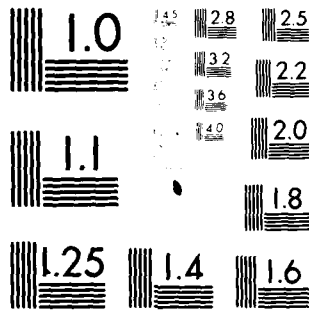
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Technical Report #64

UNBIASED L_1 ESTIMATORS
AND THEIR COVARIANCES

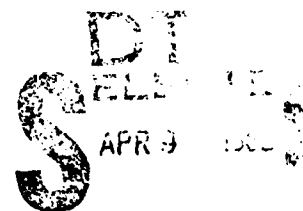
by

Book, D., Booker, J.,
Hartley, H.O., and Sielken, R.L. Jr.

Texas A&M University
Office of Naval Research
Contract N00014-78-C-0426
Project NR047-179

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ATTACHMENT I

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AND THEIR COVARIANCES

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Book, D., Booker, J.,
Hartley, H.O., and Sielken, R.L. Jr.

THEMIS OPTIMIZATION RESEARCH PROGRAM
Technical Report No. 64
June, 1980

INSTITUTE OF STATISTICS
Texas A&M University

Research conducted through the
Texas A&M Research Foundation
and sponsored by the
Office of Naval Research
Contract N00014-78-C-0426
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ATTACHMENT II

ABSTRACT

The parameters in a linear regression model can be estimated by minimizing the sum of the absolute residuals (L_1 estimation) instead of the more classical approach of minimizing the sum of squared residuals (least squares estimation). In addition to other nice properties L_1 estimators are less sensitive to outliers than least squares estimators. This paper describes a linear programming algorithm and computer program for obtaining unbiased L_1 estimators and estimates of their covariances. These estimated covariances are the new feature in this work and are an extremely important ingredient in hypothesis tests and confidence interval construction. Technical Report 65 provides an analogous treatment of L_1 estimation subject to linear constraints on the parameters.

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Unbiased L_1 Estimators and their Covariances

1. An Introduction to MRS. A.

Consider the linear regression model in the form

$$y = X\beta + \epsilon \quad (1)$$

where y is a vector of n observations, X is an $n \times p$ matrix of rank p of known constants, β is a vector of p unknown parameters and ϵ is a vector of independent random variables (noise) symmetrically distributed with mean zero and variance σ^2 . Unbiased estimation of β can be obtained under several different optimality criteria. The classical least squares approach is to estimate β by

$$\tilde{\beta} = (X^T X)^{-1} X^T Y \quad (2)$$

which has the smallest variance among the class of unbiased linear functions of y . The least squares estimator, $\tilde{\beta}$, is extremely sensitive to large values of $|\epsilon|$, outliers, particularly when the sample size, n , is small relative to p , say $n \leq 2(p+1)$. This sensitivity suggests that an optimality criteria other than minimum variance should be considered. Several authors (Barrodale (1968), Charnes and Cooper (1964), Gentle, Kennedy and Sposito (1977), Harris (1950), Harter (1974), Rice and White (1964), Taylor (1973)) have suggested that

$$\sum_{i=1}^n |y_i - X_i \beta| \quad (3)$$

should be minimized with respect to β where y_i is the i -th observation and X_i is the i -th row of X . The estimator, $\hat{\beta}$, which minimizes the sum of the absolute residuals is often called the L_1 estimator.

Since the L_1 estimate is not necessarily unique, the unbiasedness of an L_1 estimator depends upon its method of computation. Hartley and Sielken [1973] have shown how to obtain an unbiased L_1 estimator using any conventional

linear programming algorithm and an initial unbiased antisymmetrical estimator of β , say β_0 , where

$$\beta - \beta_0(\epsilon) = - [\beta - \beta_0(-\epsilon)] . \quad (4)$$

The computer program MRS. A. implements an algorithm for Minimizing the Sum of the Absolute Residuals. The algorithm uses the least squares estimator $\tilde{\beta}$ as the initial unbiased antisymmetrical estimator.

It is nice to be able to compute an L_1 estimate. The fact that the L_1 estimator can be made unbiased is the first step in understanding its properties. The second step is to estimate its covariance. Such an estimate would usually be a prerequisite for confidence intervals or hypothesis tests. In the past the absence of such a covariance estimator has made L_1 estimation less attractive. MRS. A contains a mini-Monte Carlo procedure for estimating the covariance of the L_1 estimator. This feature sets MRS. A apart from other L_1 estimation procedures.

2. Computational Procedure

The problem of minimizing the sum of the absolute residuals can be formulated as follows:

$$\min \sum_{i=1}^n r_i \quad (5)$$

subject to

$$-r_i \leq y_i - X_i \beta \leq r_i, \quad i = 1, \dots, n, \quad (6)$$

$$r_i \geq 0, \quad (7)$$

$$\beta \text{ unrestricted}, \quad (8)$$

where r_i is the i -th absolute residual. However, to insure that the resulting L_1 estimator is unbiased, the problem is reformulated following Hartley and Sielken [1973]. In particular, introducing the antisymmetrical least squares estimator, β_0 , transforms (6) to

$$-r_i \leq y_i - X_i \beta_0 - X_i (\beta - \beta_0) \leq r_i, \quad i = 1, \dots, n. \quad (9)$$

Then, using

$$\beta = \beta^{(1)} - \beta^{(2)},$$

$$\beta_0 = \beta_0^{(1)} - \beta_0^{(2)}$$

with $\beta^{(1)}, \beta^{(2)}, \beta_0^{(1)}, \beta_0^{(2)} \geq 0$ in (9) yields

$$-X_i (\beta^{(1)} + \beta_0^{(2)}) + X_i (\beta^{(2)} + \beta_0^{(1)}) - r_i \leq -y_i + X_i \beta_0$$

$$X_i (\beta^{(1)} + \beta_0^{(2)}) - X_i (\beta^{(2)} + \beta_0^{(1)}) - r_i \leq y_i - X_i \beta_0$$

or equivalently

$$\begin{aligned} -X_1 B_1 + X_1 B_2 - r_1 &\leq -y_1 + X_1 \beta_0 \\ X_1 B_1 - X_1 B_2 - r_1 &\leq y_1 - X_1 \beta_0 \end{aligned} \quad (10)$$

for $i = 1, \dots, n$ where

$$\begin{aligned} B_1 &= \beta^{(1)} + \beta_0^{(2)} \geq 0 \\ B_2 &= \beta^{(2)} + \beta_0^{(1)} \geq 0. \end{aligned} \quad (11)$$

Now, in order to compensate for any idiosyncrosies in the particular linear programming algorithm used to solve the problem in (5), (7), (10), (11), the problem is considered in two equivalent symmetrical forms, P_1 and P_2 , with MRS. A randomly selecting either P_1 or P_2 with probability $\frac{1}{2}$. The problems P_1 and P_2 in matrix notation are as follows:

$$P_1: \min \sum_{i=1}^n r_i$$

subject to

$$\begin{bmatrix} -X & X & -I \\ X & -X & -I \end{bmatrix} \begin{bmatrix} B_1 \\ B_2 \\ r \end{bmatrix} \leq \begin{bmatrix} -y + X\beta_0 \\ y - X\beta_0 \end{bmatrix}$$

$$B_1, B_2, r \geq 0;$$

$$P_2: \min \sum_{i=1}^n r_i$$

subject to

$$\begin{bmatrix} -X & X & -I \\ X & -X & -I \end{bmatrix} \begin{bmatrix} B_2 \\ B_1 \\ r \end{bmatrix} \leq \begin{bmatrix} y - X\beta_0 \\ -y + X\beta_0 \end{bmatrix},$$

$$B_1, B_2, r \geq 0.$$

After either P_1 or P_2 has been solved by MRS. A, the unbiased L_1 estimator is

$$\hat{\beta} = B_1 - B_2 + \beta_0. \quad (12)$$

If the sample size was quite large, say $n \gg (p+1)^2$, the sample could be randomly subdivided into G groups, then $\hat{\beta}_g$ estimated for each group g separately, and the covariance of $\hat{\beta}$ estimated from the sample covariance of the $\hat{\beta}_g$'s.

MRS. A estimates the covariance of $\hat{\beta}$ using a mini-Monte Carlo procedure. Conceptually a Monte-Carlo estimate would be obtained by generating several sets of n y 's, finding the L_1 estimate for each set, and computing a sample covariance. There are two difficulties with generating the y 's; namely,

- (1) β is unknown, and
- (2) σ^2 is unknown.

The first of these difficulties can be overcome by expressing the sum of absolute residuals as

$$\begin{aligned}
 & \sum_{i=1}^n |y_i - X_i \hat{\beta}| \\
 &= \sum_{i=1}^n |X_i \beta + \epsilon_i - X_i \hat{\beta}| \\
 &= \sum_{i=1}^n |\epsilon_i - X_i (\hat{\beta} - \beta)| \\
 &= \sum_{i=1}^n |\epsilon_i - X_i \delta\beta| \tag{13}
 \end{aligned}$$

using (1) and $\delta\beta = \hat{\beta} - \beta$ and noting that the covariance of the L_1 estimator of $\delta\beta$ is the same covariance of $\hat{\beta}$. Thus, only sets of n ϵ 's need be generated. To deal with the second difficulty (the unknown σ^2), note that

$$\begin{aligned}
 & \sum_{i=1}^n |\epsilon_i - X_i \delta\beta| \\
 &= \sigma \sum_{i=1}^n |\epsilon_i^* - X_i \delta\beta^*| \tag{14}
 \end{aligned}$$

where $\delta\beta^* = (\hat{\beta} - \beta) / \sigma$ and ϵ_i / σ is symmetrically distributed with mean 0 and variance 1. Furthermore,

$$\begin{aligned}
 \text{Covariance } (\hat{\beta}) &= \text{Covariance } (\delta\beta) \\
 &= \sigma^2 \text{Covariance } (\delta\beta^*). \tag{15}
 \end{aligned}$$

Thus, the mini-Monte Carlo procedure generates K sets of n ϵ^* 's; finds the K L_1 estimates $\hat{\delta\beta}^*$; and then estimates the covariance of $\hat{\beta}$ by $\hat{\sigma}^2$

times the sample covariance of the $\hat{\delta}\beta^*$.

Several possibilities for $\hat{\sigma}^2$ have been considered

$$(1) \hat{\sigma}_1^2 = \frac{1}{n} \sum_{i=1}^n (y_i - X_i \beta_0)^2 / (n-p),$$

$$(2) \hat{\sigma}_2^2 = \frac{1}{n} \sum_{i=1}^n (y_i - X_i \hat{\beta})^2 / (n-p),$$

$$(3) \hat{\sigma}_3^2 = \frac{1}{n} \sum_{i=1}^n (y_i - X_i \hat{\beta})^2 / (n-p) \\ \frac{1}{K} \sum_{k=1}^K (\epsilon_{ik}^* - X_i \hat{\delta}\beta_k^*)^2 / (n-p),$$

$$(4) \hat{\sigma}_4^2 = \left\{ \frac{1}{n} \sum_{i=1}^n |y_i - X_i \hat{\beta}| / (n-p) \right\}^2 \\ \left\{ \frac{1}{K} \sum_{k=1}^K \frac{1}{n} \sum_{i=1}^n |\epsilon_{ik}^* - X_i \hat{\delta}\beta_k^*| / (n-p) \right\}^2.$$

Of course, $\hat{\sigma}_1^2$ is the usual least squares estimator of σ^2 while $\hat{\sigma}_2^2$ has the same form but uses the L_1 estimator of β instead of the least squares estimator. The ratio estimators, $\hat{\sigma}_3^2$ and $\hat{\sigma}_4^2$ reflect the fact that the variance of the y 's is σ^2 while the variance of the ϵ^* 's is 1.

Since L_1 estimation is used to avoid certain weaknesses in least squares estimation, it seems appropriate to estimate σ^2 using the absolute residuals, as in $\hat{\sigma}_4^2$, instead of the squared residuals. Furthermore $E[|y - X\beta|]$ is often proportional to σ . In particular,

$$E[|y - X\beta|] = \sigma (2/\sqrt{\pi})^{-1}, \quad (16)$$

if ϵ has a uniform distribution;

$$E [|y - X\beta|] = \sigma (\sqrt{\pi/2})^{-1}, \quad (17)$$

if ϵ has a normal distribution; and

$$E [|y - X\beta|] = \sigma (\sqrt{2})^{-1}, \quad (18)$$

if ϵ has a double exponential distribution. Of course, if

$$E [|y - X\beta|] = C\sigma \quad (19)$$

and the ϵ^* 's are generated from the same distribution but with variance 1,

then the proportionality constant doesn't affect $\hat{\sigma}_4^2$. For these

reasons MRS. A uses $\hat{\sigma}_4^2$.

The results in (16), (17), and (18) suggest another alternative for $\hat{\sigma}$; namely,

$$\hat{\sigma}_5 = \frac{1}{C} \sum_{i=1}^n |y_i - X_i \hat{\beta}| / (n-p)$$

where C depends on the assumed distribution of ϵ . Note that $\hat{\sigma}_5$

reflects only the variability in $y_i - X_i \hat{\beta}$ whereas $\hat{\sigma}_4$ reflects not

only this variability but also the variability associated with the linear programming algorithm. Some empirical behavior of these five estimators of σ is reported in Table 1.

MRS. A allows the user to generate the ϵ^* 's from either the uniform, normal, or double exponential distributions. These distributions were selected as being representative of short, medium, and long tailed distributions respectively. These three distributions are also interesting because maximum likelihood corresponds to minimizing the maximum absolute

Table 1

Empirical Results on Five Estimators of σ^2

For $n = 20$

σ^2	$\hat{\sigma}_1^2$	$\hat{\sigma}_2^2$	$\hat{\sigma}_3^2$	$\hat{\sigma}_4^2$	$\hat{\sigma}_5^2$	ϵ distribution
25	21.14	21.66	20.36	24.24	23.45	Uniform
25	21.59	21.66	20.26	22.58	23.61	Normal
25	40.53	45.47	48.77	36.22	36.50	Double Exponential
100	111.63	119.44	109.54	110.12	113.59	Uniform
100	87.64	87.92	81.21	91.21	95.67	Normal
100	38.25	40.06	34.43	39.20	45.93	Double Exponential
400	375.31	376.33	351.39	360.95	362.25	Uniform
400	349.55	350.70	324.38	363.56	381.67	Normal
400	333.70	336.78	300.91	214.12	251.63	Double Exponential
2500	3255.33	3782.59	3551.58	3762.09	3724.80	Uniform
2500	2174.03	2181.34	2029.00	2270.24	2378.90	Normal
2500	2480.54	2907.44	2497.52	2345.71	2720.68	Double Exponential
8100	8363.82	9644.51	9027.74	8310.63	8470.63	Uniform
8100	7117.80	7139.82	6681.97	7433.47	7765.47	Normal
8100	7987.26	8063.69	7608.09	8355.12	9082.50	Double Exponential

residual if the ϵ 's are uniform, minimizing the sum of squared residuals if the ϵ 's are normal, and minimizing the sum of absolute residuals if ϵ 's are double exponential.

Since the proportionality constants $2/\sqrt{3}$, $\sqrt{\pi/2}$, and $\sqrt{2}$ or approximately 1.155, 1.253, 1.414 respectively are all nearly the same, the estimator $\hat{\sigma}_4$ is not too sensitive to the possibility that $Y - X\beta$ and ϵ^* have different distributional forms.

MRS. A allows the option of assigning weights to the residuals modifying equation (5) to

$$\min \sum_{i=1}^n W_i r_i$$

where

W_i = the weight given to the i -th residual.

If W_i are not all equal to one then the objective function in the mini-Monte Carlo study is

$$\sum_{i=1}^n W_i |\epsilon_i^* - X_i \delta \beta^*|$$

and the estimates of σ also reflect the W_i ; namely,

$$\hat{\sigma}_4 = \frac{\sum_{i=1}^n W_i |y_i - X_i \hat{\beta}| / (n-p)}{\frac{1}{K} \sum_{k=1}^K \sum_{i=1}^n W_i |\epsilon_{ik}^* - X_i \delta \hat{\beta}_k^*| / (n-p)}$$

and

$$\hat{\sigma}_5 = \frac{1}{C} \sum_{i=1}^n W_i |y_i - X_i \hat{\beta}| / (n-p)$$

3. MRS. A: User's Guide and Sample Problem

MRS. A consists of a main program and seven subroutines. The functions of these components are described in Table 2.

Input instructions are briefly documented in the program and consist of four basic card types.

The first card or card image contains the ancillary statistics for the specific problem as follows:

First Card:

<u>Card Column</u>	<u>Variable Name</u>	<u>Description</u>
1-5	NOBS	= n = number of observations (format I5; i.e., a 5 digit integer, right justified)
6-10	IP	= p = number of beta parameters (format I5)
11-15	ISAM	= K = number of samples for the mini-Monte Carlo study (format I5)
16-26	NSEED	Ten digit random number ≤ 2147483647 (format I11)
28	IWRIT1	= 1 if the main results of L_1 estimation are printed = 0 otherwise
30	IWRIT2	= 1 if the intermediate results of L_1 estimation are printed = 0 otherwise
32	IWRIT3	= 1 if the main results of the mini-Monte Carlo study are printed = 0 otherwise

<u>Card Column</u>	<u>Variable Name</u>	<u>Description</u>
34	IWRIT4	= 1 if the intermediate results of the mini-Monte Carlo study are printed = 0 otherwise
36	IWRIT5	= 1 if the inputted data are printed = 0 otherwise
38	IWRIT6	= 1 if the intermediate steps in the determination of the covariance of $\hat{\beta}$ are printed = 0 otherwise
40	IOPTN	= 1 if the ϵ^* 's are to be normally distributed = 2 if the ϵ^* 's are to be double exponentially distributed = 3 if the ϵ^* 's are to be uniformly distributed
42	IWT	= 1 if weights, W_i , are to be assigned to the residuals = 0 if residuals are not weighted

The remaining card input instructions are as follows:

<u>Card Number</u>	<u>Variable Name</u>	<u>Description</u>
Second card group:	$W_i, i=1, \dots, \text{NOBS}$	The weights assigned to the residuals [format (8F10.5) ; i.e., eight ten digit numbers with either a decimal point included or last 5 digits are assumed to be to the right of a supplied decimal point.]

<u>Card Number</u>	<u>Variable Name</u>	<u>Description</u>
Third card group:	$y_i, i=1, \dots, \text{NOBS}$	The observations [format(8F10.5)]
Fourth card group:	$X_{ij}, i=1, \dots, \text{NOBS}$ $j=1, \dots, \text{IP}$	The matrix of beta coefficients read in by rows [format (8F10.5)]

The user may also supply a title card of 80 spaces or less following the fourth card group.

The size of the problem which can be solved is limited only by the dimension statements in MRS. A. Currently these restrict the size to be 20 or less observations ($n \leq 20$), 10 or less parameters ($p \leq 10$) and 100 or less samples in the mini-Monte Carlo study ($K \leq 100$). However, expansion can easily be accomplished by increasing these dimensions in the dimension statements as documented in the program.

MRS. A is written in Fortran IV language and is compatible with Fortran G and H and WATFIV language compilers. The program uses double precision arithmetic.

MRS. A has been tested on several problems on an AMDAHL 470 V6 and should be compatible with all IBM computers. MRS. A executes small problems such as $n = 5, p = 3, K = 6$ in less than two seconds. Problems of sizes $n = 20, p = 2, K = 30$ take up to a minute of execution time.

The sample input and sample output for a sample problem are given in Appendices A and B, respectively. The program listing is given in Appendix C.

Table 2

Components of MRS. A and Their Functions

<u>Component</u>	<u>Function</u>
MAIN	Reads data and generates output. Performs L_1 estimation. Carries out the Mini-Monte Carlo study. Determines $\hat{\sigma}$ and the estimated covariance of $\hat{\beta}$.
INVERT	Inverts an $n \times n$ matrix.
CONST	Constructs the least squares estimate of β , β_0 .
XTXINV	Calculates $(X^T X)^{-1}$ for use in forming β_0 .
RAND	Generates random uniform variable with range 0 to 1.
NORMAL	Generates a vector of normally distributed ϵ^* with mean 0 and variance 1 for use in the mini-Monte Carlo study.
DOUBLE	Generates a vector of double exponentially distributed ϵ^* with mean 0 and variance 1 for use in the mini-Monte Carlo study.
UNIFORM	Generates a vector of uniformly distributed ϵ^* with mean 0 and variance 1 for use in the mini-Monte Carlo study.

REFERENCES

- Barrodale, I. (1968). L_1 Approximations and the analysis of data. Applied Statistics, 17, 51-7.
- Charnes, A. and Cooper, W.W. (1964). Absolute deviations and constrained regressions. ONR Research Memo. 96, Carnegie-Mellon University, Pittsburgh, Pa.
- Gentle, J. E., Kennedy, W.J., Sposito, V.A. (1977). On least absolute deviations estimators. Communications in Statistics A., 6, 839-45.
- Harris, T.E. (1950). Regression using minimum absolute deviations. American Statistician, 4, 14-5.
- Harter, H.L. (1974). The method of least squares and some alternatives, I. International Statistical Review, 42, 147-74.
- Harter, H.L. (1974). The method of least squares and some alternatives, II. International Statistical Review, 42, 235-64.
- Harter, H.L. (1974). The method of least squares and some alternatives, III. International Statistical Review, 43, 1-44.
- Harter, H.L. (1974). The method of least squares and some alternatives, IV. International Statistical Review, 43, 125-90 and 273-78.
- Harter, H.L. (1974). The method of least squares and some alternatives, V. International Statistical Review, 43, 269-72.
- Harter, H.L. (1974). The method of least squares and some alternatives, VI. International Statistical Review, 44, 113-59.
- Hartley, H.O. and Sielken, R.L. (1973). Two linear programming algorithms for unbiased estimation of linear models. Journal of the American Statistical Association, 68, 639-41.
- Rice, J.R. and White, J.S. (1964). Norms for smoothing and estimation. SIAM Review, 6, 243-56.
- Taylor, L.D. (1973). Estimation by minimizing the sum of absolute errors. Frontiers of Econometrics. Academic Press, New York.

APPENDIX A. SAMPLE INPUTS

5 2 20 1872539680 0 0 0 0 0 0 1 0
4.25700 -1.98300 0.02400 -3.18000 -3.58600
1.00000 1.12000
1.00000 1.95000
1.00000 3.02000
1.00000 5.43000
1.00000 6.59000

THIS IS AN EXAMPLE PROBLEM OF MRS. A. WITH NO OPTIONAL PRINTOUTS.

5 2 20 1872E39680 1 1 1 1 1 1 1 0
4.25700 -1.98300 0.02400 -3.18000 -3.98600
1.00000 1.12000
1.00000 1.95000
1.00000 3.02000
1.00000 5.43000
1.00000 6.59000

THIS IS AN EXAMPLE PROBLEM OF MRS. A. WITH ALL OPTIONAL PRINTOUTS.

APPENDIX B. SAMPLE OUTPUTS

MRS. A :

MINIMIZES SUM OF ABSOLUTE RESIDUALS.

THIS PROGRAM ESTIMATES A LINEAR REGRESSION BY MINIMIZING THE SUM OF THE ABSOLUTE RESIDUALS - L1 ESTIMATION. IN ADDITION, A MINI-MONTE CARLO SIMULATION GENERATES AN ESTIMATED COVARIANCE MATRIX FOR THE ESTIMATED REGRESSION PARAMETERS.

(Sample output with no optional printouts)

UNBIASED ESTIMATES OF THE REGRESSION PARAMETERS ARE OBTAINED USING THE PROCEDURE DESCRIBED IN A PAPER BY H.O. HARTLEY AND R.L. SIELKEN, JR, "TWO LINEAR PROGRAMMING ALGORITHMS FOR UNBIASED ESTIMATION OF LINEAR MODELS", 1973, JASA, VOL. 68, PAGES 639-41.

THE FOLLOWING PROCEDURE DEVELOPED BY :

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ROBERT L. SIELKEN, JR.

THE SUPPORT OF THE OFFICE OF NAVAL RESEARCH IS GRATEFULLY ACKNOWLEDGED.

THIS IS AN EXAMPLE PROBLEM OF MRS. A. WITH NO OPTIONAL PRINTOUTS.

NUMBER OF OBSERVATIONS = 5
NUMBER OF PARAMETERS = 2
THE SAMPLE SIZE FOR THE MINI-MONTE CARLO STUDY = 20
USER SUPPLIED INITIAL RANDOM INTEGER: NSEED = 1872539680
THE AUXILIARY LEAST SQUARES ESTIMATE, BETA0, OF THE REGRESSION PARAMETER VECTOR, BETA
LEAST SQUARES ESTIMATE OF BETA(1) = 3.227002
LEAST SQUARES ESTIMATE OF BETA(2) = -1.159747
MRS. A'S ANSWER : THE ESTIMATE OF THE REGRESSION PARAMETER VECTOR WHICH MINIMIZES THE SUM OF THE ABSOLUTE RESIDUALS:
L1 ESTIMATE OF BETA(1) = 3.416213
L1 ESTIMATE OF BETA(2) = -1.123249
THE RESIDUALS, R(I), I=1, NOBS
2.098826
3.208877
0.000000
0.496969
0.000000
THE SUM OF THE ABSOLUTE RESIDUALS = 5.804672
THE MAXIMUM ABSOLUTE RESIDUAL = 3.208877
MAIN RESULTS OF THE MINI-MONTE CARLO STUDY
ESTIMATED VALUE OF SIGMA (SIGMA HAT 4) = 2.392353
ESTIMATED COVARIANCE OF THE REGRESSION PARAMETER VECTOR (BETA) USING THIS ESTIMATE OF SIGMA
4.433552 -0.849803
-0.849803 0.364199

MRS. A :

MINIMIZES SUM OF ABSOLUTE RESIDUALS.

THIS PROGRAM ESTIMATES A LINEAR REGRESSION BY MINIMIZING THE SUM OF THE ABSOLUTE RESIDUALS - L1 ESTIMATION. IN ADDITION, A MINI-MONTE CARLO SIMULATION GENERATES AN ESTIMATED COVARIANCE MATRIX FOR THE ESTIMATED REGRESSION PARAMETERS.

UNBIASED ESTIMATES OF THE REGRESSION PARAMETERS ARE OBTAINED USING THE PROCEDURE DESCRIBED IN A PAPER BY H.O. HARTLEY AND R.L. SIELKEN, JR, "TWO LINEAR PROGRAMMING ALGORITHMS FOR UNBIASED ESTIMATION OF LINEAR MODELS", 1973, JASA, VOL. 68, PAGES 639-41.

(Sample output with optional printouts)

THE FOLLOWING PROCEDURE DEVELOPED BY :

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INQUIRIES AND COMMENTS SHOULD BE ADDRESSED TO:
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THE SUPPORT OF THE OFFICE OF NAVAL RESEARCH IS GRATEFULLY ACKNOWLEDGED.

THIS IS AN EXAMPLE PROBLEM OF MRS. A. WITH ALL OPTIONAL PRINTOUTS.

NUMBER OF OBSERVATIONS = 5
NUMBER OF PARAMETERS = 2
THE SAMPLE SIZE FOR THE MINI-MONTE CARLO STUDY = 20
USER SUPPLIED INITIAL RANDOM INTEGER: NSEED = 1872539680
THE LINEAR REGRESSION IS $Y = X\beta + \epsilon$
WHERE
Y CONTAINS THE OBSERVATIONS,
BETA IS A VECTOR CONTAINING THE REGRESSION PARAMETERS,
THE I-TH ROW OF X CONTAINS THE COEFFICIENTS OF BETA
CORRESPONDING TO THE I-TH OBSERVATION, AND
EPSILON IS A RANDOM VARIABLE WITH MEAN ZERO AND VARIANCE
SIGMA-SQUARED REPRESENTING RANDOM VARIABILITY FROM
 $X\beta$.

THE Y VECTOR

4.25700
-1.98300
0.02400
-3.18000
-3.98600

THE X MATRIX

1.0 1.1
1.0 2.0
1.0 3.0
1.0 5.4
1.0 6.6

SUPPLEMENTAL INFORMATION FROM THE LINEAR PROGRAMMING PROBLEM DETERMINATION OF THE ESTIMATE OF THE REGRESSION PARAMETER VECTOR, BETA.

THE LINEAR PROGRAMMING PROBLEM AS IT WAS CREATED

THE OBJECTIVE FUNCTION COEFFICIENTS

C(1) = 0.00000
C(2) = 0.00000
C(3) = 0.00000
C(4) = 0.00000
C(5) = -1.00000
C(6) = -1.00000
C(7) = -1.00000
C(8) = -1.00000
C(9) = -1.00000
C(10) = 0.00000
C(11) = 0.00000
C(12) = 0.00000
C(13) = 0.00000
C(14) = 0.00000
C(15) = 0.00000
C(16) = 0.00000
C(17) = 0.00000
C(18) = 0.00000
C(19) = 0.00000

THE CONSTRAINT MATRIX A

-1.0 -1.1 1.0 1.1 -1.0 0.0 0.0 0.0 0.0 1.0 0.0 0.0 0.0 0.0 0.0 0.0

0.0	0.0	0.0														
	-1.0	-2.0	1.0	2.0	0.0	-1.0	0.0	0.0	0.0	0.0	1.0	0.0	0.0	0.0	0.0	0.0
0.0	0.0	0.0														
	-1.0	-3.0	1.0	3.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	1.0	0.0	0.0	0.0	0.0
0.0	0.0	0.0														
	-1.0	-5.4	1.0	5.4	0.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	1.0	0.0	0.0	0.0
0.0	0.0	0.0														
	-1.0	-6.6	1.0	6.6	0.0	0.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	1.0	0.0	0.0
0.0	0.0	0.0														
	1.0	1.1	-1.0	-1.1	-1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	1.0	0.0
0.0	0.0	0.0														
	1.0	2.0	-1.0	-2.0	0.0	-1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	1.0
0.0	0.0	0.0														
	1.0	3.0	-1.0	-3.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
1.0	0.0	0.0														
	1.0	5.4	-1.0	-5.4	0.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
0.0	1.0	0.0														
	1.0	6.6	-1.0	-6.6	0.0	0.0	0.0	0.0	-1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
0.0	0.0	1.0														

THE VARIABLES INITIALLY IN THE BASIS
10,11,12,13,14,15,16,17,18,19,
THE PROBLEM P2 HAS BEEN SELECTED
THE INVERSE OF XTX

0.810298 -0.168497
-0.168497 0.046521
VALUE OF RETAO TO COMPUTE RHS
3.227002 -1.159747
THE RHS, YMXB, FOR P1 OR P2

2.328914
-2.948496
0.299433
-0.109578
0.429728

THE INITIAL VALUES OF THE BASIC VARIABLES

XB(1) = 0.23289D 01
XB(2) = -0.29485D 01
XB(3) = 0.29943D 00
XB(4) = -0.10958D 00
XB(5) = 0.42973D 00
XB(6) = -0.23289D 01
XB(7) = 0.29485D 01
XB(8) = -0.29943D 00
XB(9) = 0.10958D 00
XB(10) = -0.42973D 00

THE INITIAL VALUE OF THE OBJECTIVE FUNCTION = 0.00000D 00
THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
THE 2-TH REDUCED COST = -0.00000D 00
THE 3-TH REDUCED COST = -0.00000D 00
THE 4-TH REDUCED COST = -0.00000D 00
THE 5-TH REDUCED COST = -0.10000D 01
THE 6-TH REDUCED COST = -0.10000D 01
THE 7-TH REDUCED COST = -0.10000D 01
THE 8-TH REDUCED COST = -0.10000D 01
THE 9-TH REDUCED COST = -0.10000D 01
THE 10-TH REDUCED COST = -0.00000D 00
THE 11-TH REDUCED COST = -0.00000D 00
THE 12-TH REDUCED COST = -0.00000D 00
THE 13-TH REDUCED COST = -0.00000D 00
THE 14-TH REDUCED COST = -0.00000D 00
THE 15-TH REDUCED COST = -0.00000D 00
THE 16-TH REDUCED COST = -0.00000D 00
THE 17-TH REDUCED COST = -0.00000D 00
THE 18-TH REDUCED COST = -0.00000D 00
THE 19-TH REDUCED COST = -0.00000D 00

TENTATIVELY THE 2-TH BASIC VARIABLE IS LEAVING THE BASIS
THE YR(J)'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.10000D 01
YR(2) = -0.19500D 01
YR(3) = 0.10000D 01
YR(4) = 0.19500D 01
YR(5) = 0.00000D 00
YR(6) = -0.10000D 01
YR(7) = 0.00000D 00
YR(8) = 0.00000D 00
YR(9) = 0.00000D 00
YR(10) = 0.00000D 00
YR(11) = 0.00000D 00
YR(12) = 0.00000D 00

YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.00000D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 1-TH NET PRICE TO YR(1) = -0.00000D 00
 THE RATIO OF THE 2-TH NET PRICE TO YR(2) = -0.00000D 00
 THE RATIO OF THE 6-TH NET PRICE TO YR(6) = -0.10000D 01

THE 2-TH VARIABLE IS LEAVING THE BASIS.

THE 2-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

10, 2, 12, 13, 14, 15, 16, 17, 18, 19,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.40224D 01
 XB(2) = 0.15120D 01
 XB(3) = 0.48658D 01
 XB(4) = 0.81009D 01
 XB(5) = 0.10394D 02
 XB(6) = -0.40224D 01
 XB(7) = 0.22204D-15
 XB(8) = -0.48658D 01
 XB(9) = -0.81009D 01
 XB(10) = -0.10394D 02

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.00000D 00

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.00000D 00
 THE 5-TH REDUCED COST = -0.10000D 01
 THE 6-TH REDUCED COST = -0.10000D 01
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.10000D 01
 THE 9-TH REDUCED COST = -0.10000D 01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.00000D 00
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.00000D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.00000D 00
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.00000D 00

TENTATIVELY THE 10-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.

IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.23795D 01
 YR(2) = 0.00000D 00
 YR(3) = 0.23795D 01
 YR(4) = -0.44409D-15
 YR(5) = 0.00000D 00
 YR(6) = -0.33795D 01
 YR(7) = 0.00000D 00
 YR(8) = 0.00000D 00
 YR(9) = -0.10000D 01
 YR(10) = 0.00000D 00
 YR(11) = 0.33795D 01
 YR(12) = 0.00000D 00
 YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.00000D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 1-TH NET PRICE TO YR(1) = -0.00000D 00
 THE RATIO OF THE 6-TH NET PRICE TO YR(6) = -0.29590D 00
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.10000D 01

THE 10-TH VARIABLE IS LEAVING THE BASIS.

THE 1-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

10, 2, 12, 13, 14, 15, 16, 17, 18, 1,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.58817D 01
 XB(2) = -0.72807D 00

XB(3) = 0.24689D 01
 XB(4) = 0.30525D 00
 XB(5) = 0.66613D-15
 XB(6) = -0.58817D 01
 XB(7) = 0.10080D-15
 XB(8) = -0.24689D 01
 XB(9) = -0.30525D 00
 XB(10) = 0.43682D 01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.00000D 00
 THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.00000D 00
 THE 5-TH REDUCED COST = -0.10000D 01
 THE 6-TH REDUCED COST = -0.10000D 01
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.10000D 01
 THE 9-TH REDUCED COST = -0.10000D 01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.00000D 00
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.00000D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.00000D 00
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.00000D 00

TENTATIVELY THE 6-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.00000D 00
 YR(3) = -0.18041D-15
 YR(4) = -0.44409D-15
 YR(5) = -0.10000D 01
 YR(6) = -0.11789D 01
 YR(7) = 0.00000D 00
 YR(8) = 0.00000D 00
 YR(9) = -0.17888D 00
 YR(10) = 0.00000D 00
 YR(11) = 0.11789D 01
 YR(12) = 0.00000D 00
 YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.17888D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 5-TH NET PRICE TO YR(5) = -0.10000D 01
 THE RATIO OF THE 6-TH NET PRICE TO YR(6) = -0.84826D 00
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.55904D 01

THE 6-TH VARIABLE IS LEAVING THE BASIS.
 THE 6-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

10, 2, 12, 13, 14, 6, 16, 17, 18, 1,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.22204D-15
 XB(2) = 0.34720D 00
 XB(3) = -0.13698D 01
 XB(4) = -0.94206D 00
 XB(5) = -0.15495D-14
 XB(6) = 0.49892D 01
 XB(7) = 0.99785D 01
 XB(8) = 0.13698D 01
 XB(9) = 0.94206D 00
 XB(10) = -0.27178D 01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.49892D 01
 THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = 0.13878D-15
 THE 4-TH REDUCED COST = 0.27756D-15
 THE 5-TH REDUCED COST = -0.15174D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.10000D 01

THE 9-TH REDUCED COST = -0.84826D 00
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.00000D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.84826D 00
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.15174D 00

TENTATIVELY THE 10-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.00000D 00
 YR(3) = -0.10000D 01
 YR(4) = 0.00000D 00
 YR(5) = -0.12048D 01
 YR(6) = 0.00000D 00
 YR(7) = 0.00000D 00
 YR(8) = 0.00000D 00
 YR(9) = 0.20475D 00
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00
 YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.12048D 01
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = -0.20475D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 3-TH NET PRICE TO YR(3) = 0.13878D-15
 THE RATIO OF THE 5-TH NET PRICE TO YR(5) = -0.12595D 00
 THE RATIO OF THE 19-TH NET PRICE TO YR(19) = -0.74107D 00

THE 10-TH VARIABLE IS LEAVING THE BASIS.

THE 3-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

10, 2, 12, 13, 14, 6, 16, 17, 18, 3,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.14661D-15
 XB(2) = 0.34720D 00
 XB(3) = -0.13698D 01
 XB(4) = -0.94206D 00
 XB(5) = -0.53118D-15
 XB(6) = 0.49892D 01
 XB(7) = 0.99785D 01
 XB(8) = 0.13698D 01
 XB(9) = 0.94206D 00
 XB(10) = 0.27178D 01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.49892D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = 0.26368D-15
 THE 5-TH REDUCED COST = -0.15174D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.10000D 01
 THE 9-TH REDUCED COST = -0.84826D 00
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.00000D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.84826D 00
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.15174D 00

TENTATIVELY THE 3-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.13878D-16
 YR(2) = 0.00000D 00
 YR(3) = 0.00000D 00

YR(4) = 0.000000 00
 YR(5) = -0.652650 00
 YR(6) = 0.000000 00
 YR(7) = -0.100000 01
 YR(8) = 0.000000 00
 YR(9) = -0.347350 00
 YR(10) = 0.000000 00
 YR(11) = 0.000000 00
 YR(12) = 0.000000 00
 YR(13) = 0.000000 00
 YR(14) = 0.000000 00
 YR(15) = 0.652650 00
 YR(16) = 0.000000 00
 YR(17) = 0.000000 00
 YR(18) = 0.000000 00
 YR(19) = 0.347350 00
 THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S
 THE RATIO OF THE 5-TH NET PRICE TO YR(5) = -0.232490 00
 THE RATIO OF THE 7-TH NET PRICE TO YR(7) = -0.100000 01
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.244210 01
 THE 3-TH VARIABLE IS LEAVING THE BASIS.
 THE 5-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW
 10, 2, 5, 13, 14, 6, 16, 17, 18, 3,
 THE VALUES OF THE BASIC VARIABLES ARE NOW
 XB(1) = 0.419770 01
 XB(2) = -0.364970 01
 XB(3) = 0.209880 01
 XB(4) = -0.496970 00
 XB(5) = -0.688000 15
 XB(6) = 0.320890 01
 XB(7) = 0.641780 01
 XB(8) = 0.222040 15
 XB(9) = 0.496970 00
 XB(10) = 0.189210 00
 THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.530770 01
 THE REDUCED COSTS
 THE 1-TH REDUCED COST = -0.208170 15
 THE 2-TH REDUCED COST = -0.000000 00
 THE 3-TH REDUCED COST = -0.000000 00
 THE 4-TH REDUCED COST = 0.222040 15
 THE 5-TH REDUCED COST = -0.000000 00
 THE 6-TH REDUCED COST = -0.000000 00
 THE 7-TH REDUCED COST = -0.767510 00
 THE 8-TH REDUCED COST = -0.100000 01
 THE 9-TH REDUCED COST = -0.767510 00
 THE 10-TH REDUCED COST = -0.000000 00
 THE 11-TH REDUCED COST = -0.100000 01
 THE 12-TH REDUCED COST = -0.232490 00
 THE 13-TH REDUCED COST = -0.000000 00
 THE 14-TH REDUCED COST = -0.000000 00
 THE 15-TH REDUCED COST = -0.100000 01
 THE 16-TH REDUCED COST = -0.000000 00
 THE 17-TH REDUCED COST = -0.000000 00
 THE 18-TH REDUCED COST = -0.000000 00
 THE 19-TH REDUCED COST = -0.232490 00
 TENTATIVELY THE 4-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S
 (YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)
 YR(1) = -0.138780 16
 YR(2) = 0.000000 00
 YR(3) = 0.000000 00
 YR(4) = -0.666130 15
 YR(5) = 0.000000 00
 YR(6) = 0.000000 00
 YR(7) = 0.324930 00
 YR(8) = -0.100000 01
 YR(9) = -0.675070 00
 YR(10) = 0.000000 00
 YR(11) = 0.000000 00
 YR(12) = -0.324930 00
 YR(13) = 0.000000 00
 YR(14) = 0.000000 00
 YR(15) = 0.000000 00
 YR(16) = 0.000000 00
 YR(17) = 0.000000 00
 YR(18) = 0.000000 00
 YR(19) = 0.675070 00
 THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S
 THE RATIO OF THE 8-TH NET PRICE TO YR(8) = -0.100000 01

THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.11369D 01
 THE RATIO OF THE 12-TH NET PRICE TO YR(12) = -0.71552D 00
 THE 4-TH VARIABLE IS LEAVING THE BASIS.
 THE 12-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

10, 2, 5, 12, 14, 6, 16, 17, 18, 3,
 THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.88846D 01
 XB(2) = -0.46492D 00
 XB(3) = 0.44423D 01
 XB(4) = 0.15295D 01
 XB(5) = -0.86309D-15
 XB(6) = 0.12210D 01
 XB(7) = 0.24420D 01
 XB(8) = -0.15295D 01
 XB(9) = 0.55511D-16
 XB(10) = -0.26341D 01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.56633D 01
 THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.12490D-15
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = 0.66613D-15
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.28448D 00
 THE 9-TH REDUCED COST = -0.28448D 00
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.71552D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.71552D 00

TENTATIVELY THE 10-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.10000D 01
 YR(2) = 0.00000D 00
 YR(3) = 0.00000D 00
 YR(4) = -0.35527D-14
 YR(5) = 0.00000D 00
 YR(6) = 0.00000D 00
 YR(7) = 0.00000D 00
 YR(8) = -0.56810D 01
 YR(9) = -0.46810D 01
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00
 YR(13) = 0.56810D 01
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.46810D 01

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 1-TH NET PRICE TO YR(1) = -0.12490D-15
 THE RATIO OF THE 8-TH NET PRICE TO YR(8) = -0.50076D-01
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.60773D-01

THE 10-TH VARIABLE IS LEAVING THE BASIS.
 THE 1-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

10, 2, 5, 12, 14, 6, 16, 17, 18, 1,
 THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.88846D 01
 XB(2) = -0.46492D 00
 XB(3) = 0.44423D 01
 XB(4) = 0.15295D 01
 XB(5) = -0.86309D-15
 XB(6) = 0.12210D 01
 XB(7) = 0.24420D 01
 XB(8) = -0.15295D 01
 XB(9) = 0.18956D-16
 XB(10) = 0.26341D 01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.56633D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = 0.41633D-16
 THE 4-TH REDUCED COST = 0.66613D-15
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.10000D 01
 THE 8-TH REDUCED COST = -0.28448D 00
 THE 9-TH REDUCED COST = -0.28448D 00
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.71552D 00
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.00000D 00
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.71552D 00

TENTATIVELY THE 8-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YR(J)'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS,
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.00000D 00
 YR(3) = 0.00000D 00
 YR(4) = -0.15543D-14
 YR(5) = 0.00000D 00
 YR(6) = 0.00000D 00
 YR(7) = -0.10000D 01
 YR(8) = -0.30776D 01
 YR(9) = -0.20776D 01
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00
 YR(13) = 0.30776D 01
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.00000D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.20776D 01

THE RATIO OF THE NET PRICES TO THE NEGATIVE YR(J)'S

THE RATIO OF THE 7-TH NET PRICE TO YR(7) = -0.10000D 01
 THE RATIO OF THE 8-TH NET PRICE TO YR(8) = -0.92437D-01
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.13693D 00

THE 8-TH VARIABLE IS LEAVING THE BASIS.
 THE 8-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

10, 2, 5, 12, 14, 6, 16, 8, 18, 1,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.41977D 01
 XB(2) = -0.36497D-01
 XB(3) = 0.20988D 01
 XB(4) = 0.44409D-15
 XB(5) = -0.68800D-15
 XB(6) = 0.32089D 01
 XB(7) = 0.64178D 01
 XB(8) = 0.49697D 00
 XB(9) = 0.99394D 00
 XB(10) = -0.18921D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.58047D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.12490D-15
 THE 4-TH REDUCED COST = 0.66613D-15
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.90756D 00
 THE 8-TH REDUCED COST = -0.00000D 00
 THE 9-TH REDUCED COST = -0.92437D-01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.10000D 01
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.92437D-01

THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.90756D 00
 TENTATIVELY THE 10-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S
 (YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.00000D 00
 YR(3) = -0.10000D 01
 YR(4) = 0.46629D-14
 YR(5) = 0.00000D 00
 YR(6) = 0.00000D 00
 YR(7) = -0.18459D 01
 YR(8) = 0.00000D 00
 YR(9) = 0.84594D 00
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00
 YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = 0.18459D 01
 YR(18) = 0.00000D 00
 YR(19) = -0.84594D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S
 THE RATIO OF THE 3-TH NET PRICE TO YR(3) = -0.12490D-15
 THE RATIO OF THE 7-TH NET PRICE TO YR(7) = -0.49165D 00
 THE RATIO OF THE 19-TH NET PRICE TO YR(19) = -0.10728D 01
 THE 10-TH VARIABLE IS LEAVING THE BASIS.
 THE 3-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

10, 2, 5, 12, 14, 6, 16, 8, 18, 3,
 THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.41977D 01
 XB(2) = -0.36497D-01
 XB(3) = 0.20988D 01
 XB(4) = 0.40208D-15
 XB(5) = -0.69062D-15
 XB(6) = 0.32089D 01
 XB(7) = 0.64178D 01
 XB(8) = 0.49697D 00
 XB(9) = 0.99394D 00
 XB(10) = 0.18921D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.58047D 01
 THE REDUCED COSTS

THE 1-TH REDUCED COST = 0.13878D-16
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = 0.44409D-15
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.90756D 00
 THE 8-TH REDUCED COST = -0.00000D 00
 THE 9-TH REDUCED COST = -0.92437D-01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.10000D 01
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.92437D-01
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.90756D 00

TENTATIVELY THE 2-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.00000D 00
 YR(3) = 0.00000D 00
 YR(4) = -0.10000D 01
 YR(5) = 0.00000D 00
 YR(6) = 0.00000D 00
 YR(7) = 0.28011D 00
 YR(8) = 0.00000D 00
 YR(9) = -0.28011D 00
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00

YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = 0.00000D 00
 YR(16) = 0.00000D 00
 YR(17) = -0.28011D 00
 YR(18) = 0.00000D 00
 YR(19) = 0.28011D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S
 THE RATIO OF THE 4-TH NET PRICE TO YR(4) = 0.44409D-15
 THE RATIO OF THE 9-TH NET PRICE TO YR(9) = -0.33000D 00
 THE RATIO OF THE 17-TH NET PRICE TO YR(17) = -0.33000D 00

THE 2-TH VARIABLE IS LEAVING THE BASIS.
 THE 4-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

10, 4, 5, 12, 14, 6, 16, 8, 18, 3,
 THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.41977D 01
 XB(2) = 0.36497D-01
 XB(3) = 0.20988D 01
 XB(4) = 0.26962D-15
 XB(5) = -0.70683D-15
 XB(6) = 0.32089D 01
 XB(7) = 0.64178D 01
 XB(8) = 0.49697D 00
 XB(9) = 0.99394D 00
 XB(10) = 0.18921D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.58047D 01
 THE REDUCED COSTS

THE 1-TH REDUCED COST = 0.27756D-16
 THE 2-TH REDUCED COST = -0.22204D-15
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.00000D 00
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.90756D 00
 THE 8-TH REDUCED COST = -0.00000D 00
 THE 9-TH REDUCED COST = -0.92437D-01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.10000D 01
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.92437D-01
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.90756D 00

THE CURRENT BASIC SOLUTION IS FEASIBLE AND HENCE OPTIMAL.
 THE NONZERO VARIABLES ARE AS FOLLOWS:

X(10) = 0.41977D 01
 X(4) = 0.36497D-01
 X(5) = 0.20988D 01
 X(12) = 0.26962D-15
 X(14) = -0.70683D-15
 X(6) = 0.32089D 01
 X(16) = 0.64178D 01
 X(8) = 0.49697D 00
 X(18) = 0.99394D 00
 X(3) = 0.18921D 00

THE OPTIMAL VALUE OF THE OBJECTIVE FUNCTION IS -0.58047D 01

SUPPLEMENTAL INFORMATION FROM THE MINI-MONTE CARLO STUDY

THE GENERATED N(0,1) EPSILONS SAMPLE NUMBER = 1
 -0.37319
 -0.57141
 0.04588
 0.10940
 0.29791

THE PROBLEM P2 HAS BEEN SELECTED
 THE INVERSE OF XTX

0.810298 -0.168497
 -0.168497 0.046521
 VALUE OF BETA0 TO COMPUTE RMS
 -0.594250 0.136933
 THE RMS, YHXB, FOR P1 OR P2

0.067696
 -0.244174
 0.226595
 -0.039891
 -0.010226

THE INITIAL VALUES OF THE BASIC VARIABLES

XB(1) = -0.56988D 00
 XB(2) = -0.66336D-01
 XB(3) = -0.28494D 00
 XB(4) = 0.80279D-16
 XB(5) = 0.48323D-16
 XB(6) = 0.54175D 00
 XB(7) = 0.10835D 01
 XB(8) = 0.10662D 00
 XB(9) = 0.21323D 00
 XB(10) = 0.42693D 00

THE INITIAL VALUE OF THE OBJECTIVE FUNCTION = -0.36343D 00

THE REDUCED COSTS

THE 1-TH REDUCED COST = 0.27756D-16
 THE 2-TH REDUCED COST = -0.22204D-15
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.00000D 00
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.90756D 00
 THE 8-TH REDUCED COST = -0.00000D 00
 THE 9-TH REDUCED COST = -0.92437D-01
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.00000D 00
 THE 13-TH REDUCED COST = -0.10000D 01
 THE 14-TH REDUCED COST = -0.00000D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.92437D-01
 THE 18-TH REDUCED COST = -0.00000D 00
 THE 19-TH REDUCED COST = -0.90756D 00

TENTATIVELY THE 1-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YR(J)'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.

IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.00000D 00
 YR(2) = 0.22204D-15
 YR(3) = 0.00000D 00
 YR(4) = 0.00000D 00
 YR(5) = 0.00000D 00
 YR(6) = 0.00000D 00
 YR(7) = -0.30644D 01
 YR(8) = 0.00000D 00
 YR(9) = 0.10644D 01
 YR(10) = 0.00000D 00
 YR(11) = 0.00000D 00
 YR(12) = 0.00000D 00
 YR(13) = 0.00000D 00
 YR(14) = 0.00000D 00
 YR(15) = -0.10000D 01
 YR(16) = 0.00000D 00
 YR(17) = 0.30644D 01
 YR(18) = 0.00000D 00
 YR(19) = -0.10644D 01

THE RATIO OF THE NET PRICES TO THE NEGATIVE YR(J)'S

THE RATIO OF THE 7-TH NET PRICE TO YR(7) = -0.29616D 00
 THE RATIO OF THE 15-TH NET PRICE TO YR(15) = -0.10000D 01
 THE RATIO OF THE 19-TH NET PRICE TO YR(19) = -0.85263D 00

THE 1-TH VARIABLE IS LEAVING THE BASIS.

THE 7-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

7, 4, 5, 12, 14, 6, 16, 8, 18, 3,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.18596D 00
 XB(2) = -0.14245D-01
 XB(3) = 0.27756D-16
 XB(4) = 0.37193D 00
 XB(5) = 0.86648D-17
 XB(6) = 0.30005D 00
 XB(7) = 0.60009D 00
 XB(8) = 0.46190D-01
 XB(9) = 0.92380D-01
 XB(10) = 0.83651D-01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.53220D 00

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.36082D-15
 THE 2-TH REDUCED COST = -0.22204D-15
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.00000D 00
 THE 5-TH REDUCED COST = -0.00000D 00

THE 6-TH REDUCED COST = -0.000000 00
 THE 7-TH REDUCED COST = -0.000000 00
 THE 8-TH REDUCED COST = -0.000000 00
 THE 9-TH REDUCED COST = -0.407680 00
 THE 10-TH REDUCED COST = -0.296160 00
 THE 11-TH REDUCED COST = -0.100000 01
 THE 12-TH REDUCED COST = -0.000000 00
 THE 13-TH REDUCED COST = -0.100000 01
 THE 14-TH REDUCED COST = -0.000000 00
 THE 15-TH REDUCED COST = -0.703840 00
 THE 16-TH REDUCED COST = -0.000000 00
 THE 17-TH REDUCED COST = -0.100000 01
 THE 18-TH REDUCED COST = -0.000000 00
 THE 19-TH REDUCED COST = -0.592320 00

TENTATIVELY THE 2-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.138780-16
 YR(2) = -0.100000 01
 YR(3) = 0.000000 00
 YR(4) = 0.000000 00
 YR(5) = 0.000000 00
 YR(6) = 0.000000 00
 YR(7) = 0.000000 00
 YR(8) = 0.000000 00
 YR(9) = 0.182820 00
 YR(10) = -0.914080-01
 YR(11) = 0.000000 00
 YR(12) = 0.000000 00
 YR(13) = 0.000000 00
 YR(14) = 0.000000 00
 YR(15) = 0.914080-01
 YR(16) = 0.000000 00
 YR(17) = 0.000000 00
 YR(18) = 0.000000 00
 YR(19) = -0.182820 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 2-TH NET PRICE TO YR(2) = -0.222040-15
 THE RATIO OF THE 10-TH NET PRICE TO YR(10) = -0.324000 01
 THE RATIO OF THE 19-TH NET PRICE TO YR(19) = -0.324000 01

THE 2-TH VARIABLE IS LEAVING THE BASIS.
 THE 2-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

7, 2, 5, 12, 14, 6, 16, 8, 18, 3,
 THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.185960 00
 XB(2) = 0.142450-01
 XB(3) = 0.383050-16
 XB(4) = 0.371930 00
 XB(5) = 0.181540-16
 XB(6) = 0.300050 00
 XB(7) = 0.600090 00
 XB(8) = 0.461900-01
 XB(9) = 0.923800-01
 XB(10) = 0.836510-01

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.532200 00
 THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.319190-15
 THE 2-TH REDUCED COST = -0.000000 00
 THE 3-TH REDUCED COST = -0.000000 00
 THE 4-TH REDUCED COST = -0.000000 00
 THE 5-TH REDUCED COST = -0.000000 00
 THE 6-TH REDUCED COST = -0.000000 00
 THE 7-TH REDUCED COST = -0.000000 00
 THE 8-TH REDUCED COST = -0.000000 00
 THE 9-TH REDUCED COST = -0.407680 00
 THE 10-TH REDUCED COST = -0.296160 00
 THE 11-TH REDUCED COST = -0.100000 01
 THE 12-TH REDUCED COST = -0.000000 00
 THE 13-TH REDUCED COST = -0.100000 01
 THE 14-TH REDUCED COST = -0.000000 00
 THE 15-TH REDUCED COST = -0.703840 00
 THE 16-TH REDUCED COST = -0.000000 00
 THE 17-TH REDUCED COST = -0.100000 01
 THE 18-TH REDUCED COST = -0.000000 00
 THE 19-TH REDUCED COST = -0.592320 00

THE CURRENT BASIC SOLUTION IS FEASIBLE AND HENCE OPTIMAL.
 THE NONZERO VARIABLES ARE AS FOLLOWS:

X(7) = 0.185960 00
 X(2) = 0.142450-01

X(5) = 0.38305D-16
X(12) = 0.37193D 00
X(14) = 0.18154D-16
X(6) = 0.30005D 00
X(16) = 0.60009D 00
X(8) = 0.46190D-01
X(18) = 0.92380D-01
X(3) = 0.83651D-01

THE OPTIMAL VALUE OF THE OBJECTIVE FUNCTION IS -0.53220D 00

SUPPLEMENTAL INFORMATION FROM THE MINI-MONTE CARLO STUDY

THE GENERATED N(0,1) EPSILONS SAMPLE NUMBER = 2
0.87619
0.59391
-0.35016
0.51585
1.51357

THE PROBLEM P1 HAS BEEN SELECTED

THE INVERSE OF XTX

0.810298 -0.168497
-0.168497 0.046521

VALUE OF BETAO TO COMPUTE RHS

0.216974 0.113998

THE RHS, YHXB, FOR P1 OR P2

-0.531540

-0.154642

0.911404

0.320132

-0.545353

THE INITIAL VALUES OF THE BASIC VARIABLES

XB(1) = 0.14477D 01

XB(2) = 0.25252D-02

XB(3) = -0.19987D-15

XB(4) = 0.28955D 01

XB(5) = -0.13871D-15

XB(6) = -0.37899D 00

XB(7) = -0.75799D 00

XB(8) = -0.86256D 00

XB(9) = -0.17251D 01

XB(10) = -0.52871D 00

THE INITIAL VALUE OF THE OBJECTIVE FUNCTION = -0.20619D 00

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.31919D-15

THE 2-TH REDUCED COST = -0.00000D 00

THE 3-TH REDUCED COST = -0.00000D 00

THE 4-TH REDUCED COST = -0.00000D 00

THE 5-TH REDUCED COST = -0.00000D 00

THE 6-TH REDUCED COST = -0.00000D 00

THE 7-TH REDUCED COST = -0.00000D 00

THE 8-TH REDUCED COST = -0.00000D 00

THE 9-TH REDUCED COST = -0.40768D 00

THE 10-TH REDUCED COST = -0.29616D 00

THE 11-TH REDUCED COST = -0.10000D 01

THE 12-TH REDUCED COST = -0.00000D 00

THE 13-TH REDUCED COST = -0.10000D 01

THE 14-TH REDUCED COST = -0.00000D 00

THE 15-TH REDUCED COST = -0.70384D 00

THE 16-TH REDUCED COST = -0.00000D 00

THE 17-TH REDUCED COST = -0.10000D 01

THE 18-TH REDUCED COST = -0.00000D 00

THE 19-TH REDUCED COST = -0.59232D 00

TENTATIVELY THE 9-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.

IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.44409D-15

YR(2) = 0.00000D 00

YR(3) = 0.00000D 00

YR(4) = 0.22204D-15

YR(5) = 0.00000D 00

YR(6) = 0.00000D 00

YR(7) = 0.00000D 00

YR(8) = 0.00000D 00

YR(9) = 0.15759D 01

YR(10) = 0.21207D 00

YR(11) = 0.00000D 00

YR(12) = 0.00000D 00

YR(13) = -0.10000D 01

YR(14) = 0.00000D 00

YR(15) = -0.21207D 00

(ETC.)

(RECONTINUED)

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = -0.41633D-16
 XB(2) = 0.56469D-01
 XB(3) = 0.73958D 00
 XB(4) = 0.34616D 01
 XB(5) = 0.18851D 00
 XB(6) = 0.17308D 01
 XB(7) = 0.26920D-16
 XB(8) = 0.94254D-01
 XB(9) = 0.14792D 01
 XB(10) = 0.38392D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCT.ON = -0.25646D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.11241D-14
 THE 2-TH REDUCED COST = -0.00000D 00
 THE 3-TH REDUCED COST = -0.00000D 00
 THE 4-TH REDUCED COST = -0.44409D-15
 THE 5-TH REDUCED COST = -0.00000D 00
 THE 6-TH REDUCED COST = -0.00000D 00
 THE 7-TH REDUCED COST = -0.00000D 00
 THE 8-TH REDUCED COST = -0.00000D 00
 THE 9-TH REDUCED COST = -0.00000D 00
 THE 10-TH REDUCED COST = -0.00000D 00
 THE 11-TH REDUCED COST = -0.10000D 01
 THE 12-TH REDUCED COST = -0.77871D 00
 THE 13-TH REDUCED COST = -0.00000D 00
 THE 14-TH REDUCED COST = -0.72129D 00
 THE 15-TH REDUCED COST = -0.10000D 01
 THE 16-TH REDUCED COST = -0.00000D 00
 THE 17-TH REDUCED COST = -0.22129D 00
 THE 18-TH REDUCED COST = -0.10000D 01
 THE 19-TH REDUCED COST = -0.27871D 00

THE CURRENT BASIC SOLUTION IS FEASIBLE AND HENCE OPTIMAL.

THE NONZERO VARIABLES ARE AS FOLLOWS:

X(7) = -0.41633D-16
 X(2) = 0.56469D-01
 X(5) = 0.73958D 00
 X(16) = 0.34616D 01
 X(13) = 0.18851D 00
 X(6) = 0.17308D 01
 X(9) = 0.26920D-16
 X(8) = 0.94254D-01
 X(10) = 0.14792D 01
 X(3) = 0.38392D 00

THE OPTIMAL VALUE OF THE OBJECTIVE FUNCTION IS -0.25646D 01

SUPPLEMENTAL INFORMATION FROM THE MINI-MONTE CARLO STUDY

THE GENERATED N(0,1) EPSILONS SAMPLE NUMBER = 20

0.41901
 -0.48994
 -1.15302
 -0.01590
 -1.66441

THE PROBLEM P2 HAS BEEN SELECTED

THE INVERSE OF XTX

0.810298 -0.168497
 -0.168497 0.046521

VALUE OF BETA0 TO COMPUTE RHS

0.178028 -0.209520

THE RHS, YMXB, FOR P1 OR P2

0.475645
 -0.259401
 -0.698296
 0.943760
 -0.461707

THE INITIAL VALUES OF THE BASIC VARIABLES

XB(1) = -0.17131D-15
 XB(2) = -0.66271D-01
 XB(3) = 0.12999D 01
 XB(4) = -0.10196D 01
 XB(5) = 0.29647D 01
 XB(6) = -0.50980D 00
 XB(7) = -0.65143D-16
 XB(8) = 0.14823D 01
 XB(9) = 0.25997D 01
 XB(10) = -0.89844D 00

THE INITIAL VALUE OF THE OBJECTIVE FUNCTION = -0.22724D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.11241D-14
 THE 2-TH REDUCED COST = -0.00000D 00

THE 3-TH REDUCED COST = -0.000000 00
 THE 4-TH REDUCED COST = -0.44409D-15
 THE 5-TH REDUCED COST = -0.000000 00
 THE 6-TH REDUCED COST = -0.000000 00
 THE 7-TH REDUCED COST = -0.000000 00
 THE 8-TH REDUCED COST = -0.000000 00
 THE 9-TH REDUCED COST = -0.000000 00
 THE 10-TH REDUCED COST = -0.000000 00
 THE 11-TH REDUCED COST = -0.100000 01
 THE 12-TH REDUCED COST = -0.77871D 00
 THE 13-TH REDUCED COST = -0.000000 00
 THE 14-TH REDUCED COST = -0.72129D 00
 THE 15-TH REDUCED COST = -0.100000 01
 THE 16-TH REDUCED COST = -0.000000 00
 THE 17-TH REDUCED COST = -0.22129D 00
 THE 18-TH REDUCED COST = -0.100000 01
 THE 19-TH REDUCED COST = -0.27871D 00

TENTATIVELY THE 4-TH BASIC VARIABLE IS LEAVING THE BASIS
 THE YR(J)'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS,
 IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = 0.51348D-15
 YR(2) = 0.000000 00
 YR(3) = 0.000000 00
 YR(4) = -0.11102D-14
 YR(5) = 0.000000 00
 YR(6) = 0.000000 00
 YR(7) = 0.000000 00
 YR(8) = 0.000000 00
 YR(9) = 0.000000 00
 YR(10) = 0.000000 00
 YR(11) = -0.100000 01
 YR(12) = 0.12997D 01
 YR(13) = 0.000000 00
 YR(14) = -0.29972D 00
 YR(15) = 0.000000 00
 YR(16) = 0.000000 00
 YR(17) = -0.12997D 01
 YR(18) = 0.000000 00
 YR(19) = 0.29972D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YR(J)'S

THE RATIO OF THE 11-TH NET PRICE TO YR(11) = -0.100000 01
 THE RATIO OF THE 14-TH NET PRICE TO YR(14) = -0.24065D 01
 THE RATIO OF THE 17-TH NET PRICE TO YR(17) = -0.17026D 00

THE 4-TH VARIABLE IS LEAVING THE BASIS.
 THE 17-TH VARIABLE IS ENTERING THE BASIS.
 THE BASIC VARIABLES ARE NOW

7, 2, 5, 17, 13, 6, 9, 8, 10, 3,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.39224D 00
 XB(2) = 0.43600D-01
 XB(3) = 0.69886D 00
 XB(4) = 0.78448D 00
 XB(5) = 0.27098D 01
 XB(6) = -0.11102D-15
 XB(7) = 0.72940D-16
 XB(8) = 0.13549D 01
 XB(9) = 0.13977D 01
 XB(10) = -0.17438D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.24460D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.99920D-15
 THE 2-TH REDUCED COST = -0.000000 00
 THE 3-TH REDUCED COST = -0.000000 00
 THE 4-TH REDUCED COST = -0.000000 00
 THE 5-TH REDUCED COST = -0.000000 00
 THE 6-TH REDUCED COST = -0.000000 00
 THE 7-TH REDUCED COST = -0.000000 00
 THE 8-TH REDUCED COST = -0.000000 00
 THE 9-TH REDUCED COST = -0.000000 00
 THE 10-TH REDUCED COST = -0.000000 00
 THE 11-TH REDUCED COST = -0.82974D 00
 THE 12-TH REDUCED COST = -0.100000 01
 THE 13-TH REDUCED COST = -0.000000 00
 THE 14-TH REDUCED COST = -0.67026D 00
 THE 15-TH REDUCED COST = -0.100000 01
 THE 16-TH REDUCED COST = -0.17026D 00
 THE 17-TH REDUCED COST = -0.000000 00
 THE 18-TH REDUCED COST = -0.100000 01
 THE 19-TH REDUCED COST = -0.32974D 00

TENTATIVELY THE 10-TH BASIC VARIABLE IS LEAVING THE BASIS

THE YRJ'S

(YR(J) = 0 IF THE J-TH VARIABLE IS IN THE BASIS.
IT CAN BE ZERO OTHERWISE TOO.)

YR(1) = -0.10000D 01
YR(2) = 0.00000D 00
YR(3) = 0.00000D 00
YR(4) = 0.66613D-15
YR(5) = 0.00000D 00
YR(6) = 0.00000D 00
YR(7) = 0.00000D 00
YR(8) = 0.00000D 00
YR(9) = 0.00000D 00
YR(10) = 0.00000D 00
YR(11) = 0.71013D 00
YR(12) = 0.18041D-15
YR(13) = 0.00000D 00
YR(14) = -0.21013D 00
YR(15) = 0.00000D 00
YR(16) = -0.71013D 00
YR(17) = 0.00000D 00
YR(18) = 0.00000D 00
YR(19) = 0.21013D 00

THE RATIO OF THE NET PRICES TO THE NEGATIVE YRJ'S

THE RATIO OF THE 1-TH NET PRICE TO YR(1) = -0.99920D-15
THE RATIO OF THE 14-TH NET PRICE TO YR(14) = -0.31897D 01
THE RATIO OF THE 16-TH NET PRICE TO YR(16) = -0.23976D 00

THE 10-TH VARIABLE IS LEAVING THE BASIS.

THE 1-TH VARIABLE IS ENTERING THE BASIS.

THE BASIC VARIABLES ARE NOW

7, 2, 5, 17, 13, 6, 9, 8, 10, 1,

THE VALUES OF THE BASIC VARIABLES ARE NOW

XB(1) = 0.39224D 00
XB(2) = 0.43600D-01
XB(3) = 0.69886D 00
XB(4) = 0.78448D 00
XB(5) = 0.27098D 01
XB(6) = -0.12796D-15
XB(7) = 0.85041D-16
XB(8) = 0.13549D 01
XB(9) = 0.13977D 01
XB(10) = 0.17438D 00

THE CURRENT VALUE OF THE OBJECTIVE FUNCTION = -0.24460D 01

THE REDUCED COSTS

THE 1-TH REDUCED COST = -0.00000D 00
THE 2-TH REDUCED COST = -0.00000D 00
THE 3-TH REDUCED COST = -0.12490D-15
THE 4-TH REDUCED COST = -0.00000D 00
THE 5-TH REDUCED COST = -0.00000D 00
THE 6-TH REDUCED COST = -0.00000D 00
THE 7-TH REDUCED COST = -0.00000D 00
THE 8-TH REDUCED COST = -0.00000D 00
THE 9-TH REDUCED COST = -0.00000D 00
THE 10-TH REDUCED COST = -0.00000D 00
THE 11-TH REDUCED COST = -0.82974D 00
THE 12-TH REDUCED COST = -0.10000D 01
THE 13-TH REDUCED COST = -0.00000D 00
THE 14-TH REDUCED COST = -0.67026D 00
THE 15-TH REDUCED COST = -0.10000D 01
THE 16-TH REDUCED COST = -0.17026D 00
THE 17-TH REDUCED COST = -0.00000D 00
THE 18-TH REDUCED COST = -0.10000D 01
THE 19-TH REDUCED COST = -0.32974D 00

THE CURRENT BASIC SOLUTION IS FEASIBLE AND HENCE OPTIMAL.

THE NONZERO VARIABLES ARE AS FOLLOWS:

X(7) = 0.39224D 00
X(2) = 0.43600D-01
X(5) = 0.69886D 00
X(17) = 0.78448D 00
X(13) = 0.27098D 01
X(6) = -0.12796D-15
X(9) = 0.85041D-16
X(8) = 0.13549D 01
X(10) = 0.13977D 01
X(1) = 0.17438D 00

THE OPTIMAL VALUE OF THE OBJECTIVE FUNCTION IS -0.24460D 01

THE AUXILIARY LEAST SQUARES ESTIMATE, BETA0, OF THE REGRESSION PARAMETER VECTOR, BETA

LEAST SQUARES ESTIMATE OF BETA(1) = 3.227002

LEAST SQUARES ESTIMATE OF BETA(2) = -1.159747

MRS. A'S ANSWER : THE ESTIMATE OF THE REGRESSION PARAMETER VECTOR WHICH MINIMIZES THE SUM OF THE ABSOLUTE RESIDUALS:

L1 ESTIMATE OF BETA(1) = 3.416213
 L1 ESTIMATE OF BETA(2) = -1.123249
 THE RESIDUALS, R(I), I=1, NOBS

2.098826
 3.208877
 0.000000
 0.496969
 0.000000

THE SUM OF THE ABSOLUTE RESIDUALS = 5.804672
 THE MAXIMUM ABSOLUTE RESIDUAL = 3.208877
 AUXILIARY RESULTS OF THE MINI-MONTE CARLO STUDY

VALUES OF DELTA BETA STAR

SAMPLE NUMBER = 1	-0.51060	0.12269
SAMPLE NUMBER = 2	0.63765	-0.02243
SAMPLE NUMBER = 3	0.94741	-0.45843
SAMPLE NUMBER = 4	0.05573	-0.14477
SAMPLE NUMBER = 5	-0.08930	-0.16773
SAMPLE NUMBER = 6	-1.10362	0.22395
SAMPLE NUMBER = 7	-0.86781	-0.04936
SAMPLE NUMBER = 8	-1.24035	0.54421
SAMPLE NUMBER = 9	0.44742	0.24662
SAMPLE NUMBER = 10	-1.03640	0.36600
SAMPLE NUMBER = 11	0.83504	0.01230
SAMPLE NUMBER = 12	0.32935	-0.04047
SAMPLE NUMBER = 13	-1.27719	0.27656
SAMPLE NUMBER = 14	-0.58869	0.33108
SAMPLE NUMBER = 15	-0.04885	0.14684
SAMPLE NUMBER = 16	0.98576	-0.03155
SAMPLE NUMBER = 17	-1.01343	0.04536
SAMPLE NUMBER = 18	1.70310	-0.28731
SAMPLE NUMBER = 19	-0.37732	-0.04003
SAMPLE NUMBER = 20	0.00365	-0.25312

ESTIMATED COVARIANCE OF DELTA BETA STAR

0.774642 -0.148480
 -0.148480 0.063634

SUM OF THE OPTIMAL OBJECTIVE FUNCTIONS OVER ALL SAMPLES = 48.526882

AUXILIARY RESULT: SIGMA HAT 5 = 2.425026

MAIN RESULTS OF THE MINI-MONTE CARLO STUDY

ESTIMATED VALUE OF SIGMA (SIGMA HAT 4) = 2.392353

ESTIMATED COVARIANCE OF THE REGRESSION PARAMETER VECTOR (BETA) USING THIS ESTIMATE OF SIGMA:
 4.433552 -0.849803

?

APPENDIX C. PROGRAM LISTING

MRSA :

MINIMIZES SUM OF ABSOLUTE RESIDUALS.

THIS PROGRAM ESTIMATES A LINEAR REGRESSION BY MINIMIZING THE SUM OF THE ABSOLUTE RESIDUALS - L1 ESTIMATION.

L1 ESTIMATION -

THIS PROGRAM USES THE DUAL SIMPLEX ALGORITHM TO IMPLEMENT THE L1 ESTIMATION PROCEDURE OUTLINED IN THE PAPER: HARTLEY AND SIEGEL, "DUAL LINEAR PROGRAMMING ALGORITHMS FOR UNBIASED ESTIMATION OF LINEAR MODELS", 1973, JASA, VOL 68, NO.343, PAGES 637-641.

THE VARIANCE OF THIS UNBIASED L1 ESTIMATOR IS ESTIMATED USING A MINI MONTE CARLO APPROACH.

THE SUM OF THE ABSOLUTE RESIDUALS, R(I), IS MINIMIZED SUBJECT TO:
-R, LE, Y-X, LE, R
R UNRESTRICTED,
R, GE, 0.

WHERE:

Y = A VECTOR OF N OBS OBSERVATIONS
X = A N OBS BY IP MATRIX OF CONSTANTS
B = AN IP BY 1 VECTOR OF UNKNOWN PARAMETERS.

THE DUAL SIMPLEX ALGORITHM . . .

THE LINEAR PROGRAMMING PROBLEM IS PUT INTO THE FORM

MAX CX
SUBJECT TO

AX = BRHS

X GREATER THAN OR EQUAL TO 0

WHERE

BRHS IS A COLUMN OF CONSTANTS

A IS AN M-BY-N MATRIX OF CONSTANTS

THE AUGMENTED VERSIONS OF A, BRHS, B, X, AND Y ARE REFERRED TO AS A1, BRHS1, B1, X1, AND Y1 RESPECTIVELY.

THE FOLLOWING PROCEDURE WAS DEVELOPED BY :

D.H. BOOK

J.R. BOOGER

H.O. HARTLEY

R. L. SIEGEL, JR.

INSTITUTE OF STATISTICS

TEXAS A&M UNIVERSITY

COLLEGE STATION, TEXAS 77843

INQUIRIES AND COMMENTS SHOULD BY ADDRESSED TO :

FOREST L. SIEGEL, JR.

THE SUPPORT OF THE OFFICE OF NAVAL RESEARCH IS GRATEFULLY

ACKNOWLEDGED.

IMPLICIT REAL*8 (A-H,O-Z)

DIMENSION C(80), ERHS(141), A1(41,80), IBASE(40)

DIMENSION DELTA(10), MT(20), TITLE(20)

DIMENSION X(141), Y(141), RECDOS(80), ISAT(80), YR(80)

DIMENSION XBT(180), ESTAR(20), DELTA(10,100), CAPB1(10), CAPB2(10)

DIMENSION IEX(80), VAR(10,10), VARDS(10,10), SUM(10)

DIMENSION X(20,10), Y(20), DB1(10), DB2(10), DBETA(10)

DIMENSION BHAT(10), BETAO(10)

COMMON X1X1,X1Y(41,41)

COMMON IURITE/IURIT1,IURIT2,IURIT3,IURIT4,IURIT5,IURIT6

COMMON/NEEVO/NEEVO

COMMON/ICONS1/ITERA,ITP,MP1,IPAP2

DOUBLE PRECISION BIINV(41,41)

THE DIMENSIONED ARRAYS HAVE THE FOLLOWING DIMENSIONS:

C(N), ERHS(M+1), A1(M+1,N), IBASE(N), ISAT(N)

BIINV(M+1,M+1), XBT(M+1), Y(M+1), RECDOS(N), YR(N)

XBT(N), ESTAR(NDBS), DELTA(IP,ISAM), CAPB1(IP), CAPB2(IP)

SUM(IP), MT(NDBS)

VARDS(IP,IP), VAR(IP,IP), DBX(N)

X(NDBS,IP), Y(NDBS), DB1(IP), DB2(IP), BETAO(IP), BHAT(IP)

DELTA (IP)

WHERE:

IP = THE NUMBER OF PARAMETERS

NDBS = THE NUMBER OF OBSERVATIONS

ISAM = THE NUMBER OF MONTE CARLO SAMPLES

M = THE NUMBER OF CONSTRAINTS = 2*NDBS

N = THE NUMBER OF VARIABLES = 2*IP + 3*NDBS

THE FOLLOWING 'TOLERANCES' ARE USED IN THE ALGORITHM. THEY WOULD BE ZERO EXCEPT FOR THE NUMERICAL INACCURACY OF THE COMPUTER.

TOLR1 : IF THE MAX REDUCED COST IS LESS THAN

TOLR1 THEN ALL REDUCED COSTS ARE CONSIDERED TO BE

NON-POSITIVE.

TOLR2 : ANY COMPONENT Y(I,J) >OR= TOLR2 IS

CONSIDERED NON-NEGATIVE

TOLR3 : IF A SINGLE VARIABLE IS, OR = TOLR3, IT IS

CONSIDERED TO BE NON-NEGATIVE

TOLR1=1.0D-07

TOLR2=-1.0D-07

TOLR3=-1.0D-07

THE INPUT - CARD NUMBER ONE.

NDBS = NUMBER OF OBSERVATIONS,

IP = NUMBER OF PARAMETERS,

ISAM = NUMBER OF SAMPLES,

NEEVO = 10 DIGIT RANDOM NUMBER LESS THAN 2147483647,

AND IURIT1, IURIT2, IURIT3, IURIT4, IURIT5, IURIT6 AND IORIN

WHERE:

THE IURIT VARIABLES INDICATE WHETHER OR NOT A

WRITE STATEMENT IS TO BE PRINTED.

IF IURIT1=1, THEN PRINTING OF THE MAIN DUAL SIMPLEX OCCURS.

ACKNOWLEDGED.

IMPLICIT REAL*8 (A-H,O-Z)
 DIMENSION C(80),BRHS(41),A1(41,80),INBASE(40)
 DIMENSION DELTA(10),WT(20),TITLE(20)
 DIMENSION XBI(41),YI(41),REDCOS(80),ISTAT(80),YR(80)
 DIMENSION XBINT(80),ESTAR(20),DELTA(10,100),CAPB1(10),CAPB2(10)
 DIMENSION DBX(80),VARB(10,10),VAROBS(10,10),SUM(10)
 DIMENSION X(20,10),Y(20),DB1(10),DB2(10),DBETA0(10)
 DIMENSION BHAT(10),BETA0(10)
 COMMON/XTXIN/XTX(41,41)
 COMMON/IKRIT/IKRIT1,IKRIT2,IKRIT3,IKRIT4,IKRIT5,IKRIT6
 COMMON/NEED/NEED
 COMMON/ICONST/ITERA,ITIP,MP1,IP1P2
 DOUBLE PRECISION B1INV(41,41)

THE DIMENSIONED ARRAYS HAVE THE FOLLOWING DIMENSIONS:
 C(N),BRHS(M1),A1(M1,N),INBASE(M),ISTAT(N)
 B1INV(M1,M1),XB1(M1),YI(M1),REDCOS(N),YR(N)
 XBINT(N),ESTAR(NOBS),DELTA(IP,ISAM),CAPB1(IP),CAPB2(IP)
 SUM(IP),WT(NOBS)
 VARDES(IP,IP),VARB(IP,IP),DBX(N)
 X(NOBS,IP),Y(NOBS),DB1(IP),DB2(IP),BETA0(IP),BHAT(IP)
 DELTA (IP)

WHERE:
 IP = THE NUMBER OF PARAMETERS
 NOBS = THE NUMBER OF OBSERVATIONS
 ISAM = THE NUMBER OF MONTE CARLO SAMPLES
 M = THE NUMBER OF CONSTRAINTS = 2*IP
 N = THE NUMBER OF VARIABLES = 2*IP + 3*NOBS

THE FOLLOWING 'TOLERANCES' ARE USED IN THE ALGORITHM.
 THEY WOULD BE ZERO EXCEPT FOR THE NUMERICAL INACCURACY OF THE
 COMPUTER
 TOLR1 : IF THE MAX REDUCED COST IS LESS THAN
 TOLR1 THEN ALL REDUCED COSTS ARE CONSIDERED TO BE
 NON-POSITIVE.
 TOLR2 : ANY COMPONENT Y(I,J) >OR = TOLR2 IS
 CONSIDERED NON-NEGATIVE
 TOLR3 : IF A BASIC VARIABLE IS < OR = TOLR3, IT IS
 CONSIDERED TO BE NON-NEGATIVE

TOLR1=1.0D-07
 TOLR2=-1.0D-07
 TOLR3=-1.0D-07

THE INPUT - CARD NUMBER ONE.

NOBS = NUMBER OF OBSERVATIONS,
 IP = NUMBER OF PARAMETERS,
 ISAM = NUMBER OF SAMPLES,
 NEED = 10 DIGIT RANDOM NUMBER LESS THAN 2147483647,
 AND IKRIT1, IKRIT2, IKRIT3, IKRIT4, IKRIT5, IKRIT6; AND 10PTN
 WHERE: THE IKRIT VARIABLES INDICATE WHETHER OR NOT A
 WRITE STATEMENT IS TO BE PRINTED.
 IF IKRIT1=1, THEN PRINTING OF THE MAIN DUAL SIMPLEX OCCURS.

MINIMIZES SUM OF ABSOLUTE RESIDUALS.

THIS PROGRAM ESTIMATES A LINEAR REGRESSION BY MINIMIZING
 THE SUM OF THE ABSOLUTE RESIDUALS - L1 ESTIMATION.

L1 ESTIMATION -

THIS PROGRAM USES THE DUAL SIMPLEX ALGORITHM TO IMPLEMENT
 THE L1 ESTIMATION PROCEDURE OUTLINED IN THE PAPER:
 HARTLEY AND STEINEN "THE LINEAR PROGRAMMING ALGORITHMS
 FOR UNBIASED ESTIMATION OF LINEAR MODELS",
 1973, JASA, VOL 68, NO.343, PAGES 639-641.

THE VARIANCE OF THIS UNBIASED L1 ESTIMATOR IS ESTIMATED
 USING A MINI MONTE CARLO APPROACH.
 THE SUM OF THE ABSOLUTE RESIDUALS, R(I), IS MINIMIZED SUBJECT TO
 -R .LE. Y-XS .LE. R
 R UNRESTRICTED,
 P .GE. 0.

WHERE:
 Y = A VECTOR OF NOBS OBSERVATIONS
 X = A NOBS BY IP MATRIX OF CONSTANTS
 B = AN IP BY 1 VECTOR OF UNKNOWN PARAMETERS.

THE DUAL SIMPLEX ALGORITHM . . .

THE LINEAR PROGRAMMING PROBLEM IS PUT INTO THE FORM
 MAX CX
 SUBJECT TO
 AX = BRHS
 X GREATER THAN OR EQUAL TO 0
 WHERE
 BRHS IS A COLUMN OF CONSTANTS
 A IS AN M-BY-N MATRIX OF CONSTANTS

THE PRESENTED VERSIONS OF ABRHS,B,AX,AND Y ARE REFERRED TO AS
 ALTERNATIVE1,XBI,AND YI RESPECTIVELY.

THE FOLLOWING PROCEDURE WAS DEVELOPED BY :
 D.N. BOOK
 J.R. BOOHER
 H.O. HARTLEY
 R.L. STEINEN, JR.
 INSTITUTE OF STATISTICS
 TEXAS A&M UNIVERSITY
 COLLEGE STATION, TEXAS 77843

INQUIRIES AND COMMENTS SHOULD BY ADDRESSED TO :
 ROBERT L. STEINEN, JR.

THE SUPPORT OF THE OFFICE OF NAVAL RESEARCH IS GRATEFULLY

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C      IF IWRIT2=1, THEN PRINTING OF INTERMEDIATE STEPS OCCURS.
C      IF IWRIT3=1, THEN PRINTING OF THE MAIN DUAL SIMPLEX STEPS
C      IN THE MONTE CARLO SAMPLES OCCURS.
C      IF IWRIT4=1, THEN PRINTING OF THE INTERMEDIATE MONTE CARLO
C      MONTE CARLO STEPS OCCURS.
C      IF IWRIT5=1, THEN PRINTING OF THE INPUTTED VALUES OF
C      X AND Y OCCURS.
C      IF IWRIT6=1, THEN PRINTING OF THE INTERMEDIATE CALCULATIONS
C      FOR THE ESTIMATED VARIANCE OF B OCCURS.
C      IF ANY IWRIT VARIABLE = 0, THEN NO PRINTING OCCURS.

C      IOPTN = 1, IF THE GENERATED EPSILONS IN THE MINI-MONTE CARLO
C      STUDY ARE DISTRIBUTED AS NORMAL RANDOM VARIABLES.
C      IOPTN = 2, IF THE GENERATED EPSILONS IN THE MINI-MONTE CARLO
C      STUDY ARE DISTRIBUTED AS DOUBLE EXPONENTIALS.
C      IOPTN = 3, IF THE GENERATED EPSILONS IN THE MINI-MONTE CARLO
C      STUDY ARE DISTRIBUTED AS UNIFORM RANDOM VARIABLES.

C      INT = 1, IF THE RESIDUALS ARE ASSIGNED WEIGHT COEFFICIENTS.
C      = 0, IF NO WEIGHTS ARE ASSIGNED TO RESIDUALS.

C      READ(5,100) NOBS,IP,ISAM,NSD,1WRIT1,1WRIT2,1WRIT3,1WRIT4,1WRIT5,
100  1WRIT6,IOPTN,INT
C      FORMAT(3I5,11I,6I2,12I2)
C      ITERA=0
C      SUMORJ=0.00

C      PRINTING OF HEADINGS AND INFORMATION

C      WRITE(6,2321)
2321  FORMAT(1H1,19X,'MRS. A :',19X,
C      *MINIMIZES SUM OF ABSOLUTE RESIDUALS.',19X,
C      * THIS PROGRAM ESTIMATES A LINEAR REGRESSION BY MINIMIZING',
19X,'THE SUM OF THE ABSOLUTE RESIDUALS - LI ESTIMATION.',19X,
19X,'IN ADDITION, A MINI-MONTE CARLO SIMULATION GENERATES AN',
19X,'ESTIMATED COVARIANCE MATRIX FOR THE ESTIMATED',19X,
C      *REGRESSION PARAMETERS.',19X, UNBIASED ESTIMATES OF THE',
C      *REGRESSION',19X,'PARAMETERS ARE OBTAINED USING THE PROCEDURE',
19X,'DESCRIBED IN A PAPER BY H.O. HARTLEY AND R.L. SIELKEN, JR.',
19X,'OF LINEAR MODELS', 1973, JASA, VOL. 68, PAGES 639-41.',19X,
19X,'THE FOLLOWING PROCEDURE DEVELOPED BY',19X,15X,
19X,'BOOK',19X,'J.P. FODER',19X,'H.O. HARTLEY',19X,
19X,'R.L. SIELKEN, JR.',19X,'INSTITUTE OF STATISTICS',19X,
19X,'TEXAS A & M UNIVERSITY',19X,'COLLEGE STATION, TEXAS 77843',
19X,'INQUIRIES AND COMMENTS SHOULD BE ADDRESSED TO:',19X,
19X,'ROBERT L. SIELKEN, JR.',19X,'THE SUPPORT OF THE',
C      *OFFICE OF NAVAL RESEARCH IS GRATEFULLY ACKNOWLEDGED.')

C      INPUT - SECOND CARD GROUP, READ ONLY IF RESIDUALS ARE
C      ASSIGNED WEIGHTS (INT=1).

C      USER SUPPLIED WEIGHT COEFFICIENTS TO ASSIGN WEIGHTS TO THE
C      RESIDUALS (OBJECTIVE FUNCTION WEIGHTS) INPUTED HERE, 16 PER
C      CARD IN F5.1 FORMAT.

C      IF (INT.EQ.1) READ(5,2329) (WT(J),J=1,NOBS)
2329  FORMAT(8F10.5)

C      PRINTING OF THE INPUTTED QUANTITIES.

C      WRITE(6,2300) NOBS,IP,ISAM,NSD
2300  FORMAT(1H0,5X,'NUMBER OF OBSERVATIONS = ',15,/,
16X,'NUMBER OF PARAMETERS = ',15,/,
16X,'THE SAMPLE SIZE FOR THE MINI-MONTE CARLO STUDY = ',15,/,

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CAPR2(L)=XBINT(L)
CAPR1(L)=XBINT(L+1P)
2702 BHAT(L)=CAPR1(L)-CAPR2(L)+BETA0(L)
GO TO 2701
2700 DO 2040 L=1,IP
CAPR3(L)=XBINT(L)
CAPR2(L)=XBINT(L+1P)
2040 BHAT(L)=CAPR1(L)-CAPR2(L)+BETA0(L)
2701 CONTINUE
IF (IURIT1.EQ.1.OR.IURIT2.EQ.1.OR.IURIT3.EQ.1.OR.IURIT4.EQ.1.OR.
* IURIT5.EQ.1.OR.IURIT6.EQ.1.) WRITE(6,858)
858 FORMAT(1H1)
WRITE(6,2326)
2326 FORMAT(1H0,5X,'THE AUXILIARY LEAST SQUARES ESTIMATE, BETA0,',
* ' OF THE REGRESSION PARAMETER VECTOR, BETA')
GO 3007 I=1,IP
3007 WRITE(6,3008) I,BETA0(I)
3008 FORMAT(1H ,12X,'LEAST SQUARES ESTIMATE OF BETA('',13,'') = ',F14.6)
WRITE(6,3009)
3009 FORMAT(1H0,5X,'MRS. A',1H,'S ANSWER ',5X,
* 'THE ESTIMATE OF THE REGRESSION PARAMETER VECTOR WHICH ',
* 'MINIMIZES THE SUM OF THE ABSOLUTE RESIDUALS: ')
DO 3010 I=1,IP
3010 WRITE(6,3011) I,BHAT(I)
3011 FORMAT(1H ,12X,'L1 ESTIMATE OF BETA('',13,'') = ',F14.6)
C
C
C PRINTOUT OF RESIDUALS, R(I), I=1,NORS
WRITE(6,3019)
3019 FORMAT(1H0,5X,'THE RESIDUALS, R(I),I=1,NORS')
DO 2067 L=1,NORS
LL=L+2*IP
2067 WRITE(6,3018) XBINT(LL)
3018 FORMAT(1H0,15X,F14.6)
WRITE(6,3118) SUMRES
3118 FORMAT(1H0,5X,'THE SUM OF THE ABSOLUTE RESIDUALS = ',F14.6)
RESMAX=0.00
DO 3119 L=1,NORS
LL=L+2*IP
IF (RESMAX.LT.DABS(XBINT(LL))) RESMAX=DABS(XBINT(LL))
3119 CONTINUE
WRITE(6,3120) RESMAX
3120 FORMAT(1H0,5X,'THE MAXIMUM ABSOLUTE RESIDUAL = ',F14.6)
C
C
C FORM THE VARIANCE OF DELTA BETA STAR
THESE CALCULATIONS USE THE EPSILON STARS.
SAH = ISAH
P = IP
DO 2062 J=1,IP
SMH=0.00
DO 2064 L=1,ISAH
2064 SMH = SMH + DELR(J,L)
2062 SUM(J)=SMH
DO 2061 I=1,IP
DO 2061 J=1,IP
SUMSR=0.00
SOSUM=0.00
SOSUM = SOSUM + SUM(J)*SUM(I)
DO 2060 L=1,ISAH

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SUMRES=0.00
DO 2022 NM=1,NORS
LH=NM+2*IP
2022 SUMRES=SUMRES+DABS(XBINT(LNM))
2020 IF (ITERA.EQ.1) GO TO 200
403 WRITE(6,530)
530 FORMAT(1H0,5X,'ALL OF THE YRJS ARE NON-NEGATIVE.',11X,'HENCE THE DUAL
* PROBLEM HAS NO FEASIBLE SOLUTION.')
IF (IURIT2.EQ.1) WRITE(6,850)
GO TO 999
999 CONTINUE
402 WRITE(6,1010)
1010 FORMAT(1H0,5X,'ALL OF THE REDUCED COSTS ARE NOT NON-POSITIVE.',11X,'HENCE THE DUAL
* PROBLEM HAS NO FEASIBLE SOLUTION.')
WRITE(6,850)
GO TO 999
999 CONTINUE
200
C
C TEST FOR COMPLETION OF ALL SAMPLES
C
C IF (ITERA.EQ.ISAH) GO TO 2999
C
C INCREMENT ITERA (THE SAMPLE INDEX COUNTER) BY 1.
C
C ITERA=ITERA+1
C
C CALL SUBROUTINE NORMAL OR DOUBLE OR UNIFORM TO GENERATE A SET
C OF RANDOM VARIABLES, EPSILONS, DISTRIBUTED AS NORMAL OR
C DOUBLE EXPONENTIAL OF UNIFORM, RESPECTIVELY, WITH MEAN
C ZERO AND VARIANCE ONE.
C
C IF (IOPTN.EQ.1) CALL N 'AL(NORS,ESTAR)
C IF (IOPTN.EQ.2) CALL DBLE(NORS,ESTAR)
C IF (IOPTN.EQ.3) CALL UNIF(NORS,ESTAR)
C IF (IURIT3.EQ.1) WRITE(6,3000) ITERA
3000 FORMAT(1H0,5X,'SUPPLEMENTAL INFORMATION FROM THE MINI-MONTE CARLO
* STUDY ',11X,'THE GENERATOR',12X,'EPSILONS',
* 'SAMPLE NUMBER = ',15)
IF (IURIT3.EQ.0) GO TO 3333
DO 3002 K=1,NORS
3002 WRITE (6,3001) ESTAR(K)
3001 FORMAT(1H ,15X,F12.5)
3333 CONTINUE
C
C CALL SUBROUTINE CONST TO CONSTRUCT XB1 FOR THIS SAMPLE
C
C IURIT1=0
C IURIT2=0
C IF (IURIT3.EQ.1) IURIT1=1
C IF (IURIT4.EQ.1) IURIT2=1
C CALL CONST(ESTAR,X,IP,NORS,XB1,DRETA0,B1INV,A1)
C GO TO 700
2999 CONTINUE
C
C ALL SAMPLES COMPLETE
C
C FORMS THE L1 ESTIMATE.
C
C IF (IURIT2.EQ.1) GO TO 2700
C GO 2702 L=1,IP

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C THIS SUBROUTINE INVERTS AN N BY N MATRIX
C
      IMPLICIT REAL*8 (A-H,O-Z)
      DIMENSION A(41,41)
      DO 51 I=1,N
        IF(A(I,I).EQ.0.D0) A(I,I)=0.1D-30
        A(I,I)=1.0D0/A(I,I)
        DO 52 J=I+1,N
          IF(J-I).53.52.53
            A(I,J)=A(J,I)*A(I,I)
          CONTINUE
          DO 51 J=I+1,N
            IF(J-I).54.51.54
              DO 56 K=J+1,N
                IF(K-J).55.56.55
                  A(J,K)=A(K,J)-A(J,I)*A(I,K)
                CONTINUE
                A(J,I)=-A(J,I)*A(I,I)
              CONTINUE
            RETURN
          END
        SURROUTINE CONST(ORS,X,IP,NORS,XB1,BETA0,BIINV,A1)
      C
      C THIS IS SUBROUTINE CONSTRUCT. THIS SUBROUTINE:
      C
      C 1. RANDOMLY SELECTS BETWEEN LINEAR PROBLEMS
      C P1 AND P2 EACH WITH A PROBABILITY OF ONE-HALF.
      C
      C 2. COMPUTES THE VALUE OF BETA0, THE LEAST
      C SQUARES ESTIMATOR WHICH SERVES AS THE INITIAL SYMMETRIC
      C ESTIMATOR.
      C
      C 3. COMPUTES THE INITIAL BASIC VARIABLES,
      C XB1, FOR THE DUAL SIMPLEX ALGORITHM.
      C
      IMPLICIT REAL*8 (A-H,O-Z)
      DIMENSION ORS(20),X(20,10),XB1(41),BETA0(10),BIINV(41,H+1)
      DIMENSION A1(41,90),YMXB(20),XTY(10)
      COMMON/XTXH/XTX(41,41)
      COMMON/IWRITE/IWRITE1,IWRITE2,IWRITE3,IWRITE4,IWRITE5,IWRITE6
      COMMON/NFEED/NFEED
      COMMON/ICONST/ITERA,ITIP,MFI,IP1P2
      C
      C DIMENSIONS OF ABOVE ARRAYS ARE . . .
      C ORS(NORS), X(NORS,IP)*XB1(MFI),BETA0(IP),BIINV(MFI,H+1)
      C A1(H+1,N),YMXB(NORS),XTX(IP,IP),XTY(IP)
      C
      C THE RANDOM SELECTION OF PROBLEMS P1 AND P2, EACH WITH
      C PROBABILITY OF ONE-HALF USING SUBROUTINE RAND.
      C
      C IF IP1P2 = 1 THEN P1 IS USED,
      C IF IP1P2 = 2 THEN P2 IS USED.
      C
      NFEED=NFEED + ITERA*70968
      CALL RAND(U)
      IP1P2=1
      IF(U.GT.5.D-1) IP1P2=2
      IF(IWRITE1.EQ.1) WRITE(5,50) IP1P2
      C
      C 50 FORMAT('H0,' THE PROBLEM P',I1,' HAS BEEN SELECTED')
      C

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C      CALCULATION OF THE LEAST SQUARES ESTIMATE OF BETA
C      USING SUBROUTINES XTXINV AND INVER.
C
      CALL XTXINV(X,IP,NORS)
      DO 31 I=1,IP
      SUM=0.00
      DO 30 J=1,NORS
      SUM = SUM + X(J,I)*ORS(J)
      XTI(I) = SUM
      31 CONTINUE
      DO 41 J=1,IP
      SUM =0.00
      DO 40 I=1,IP
      SUM = SUM + XTX(J,I)*XTY(I)
      BETAO(J) = SUM
      40 CONTINUE
      41 CONTINUE
      55 IF(IWRIT1.EQ.1) WRITE(6,55)
      56 FORMAT(1H, ' VALUE OF BETAO TO COMPUTE RHS')
      IF(IWRIT1.EQ.1) WRITE(6,56) (BETAO(J),J=1,IP)
      56 FORMAT(1H, '9X,8F14.6)

C      COMPUTATION OF THE INITIAL VALUES OF THE BASIC VARIABLES,
C      XBI, TO BE USED IN THE DUAL SIMPLEX ALGORITHM.
C
      DO 11 I=1,NORS
      SUM=ORS(I)
      DO 12 J=1,IP
      SUM = SUM - X(I,J)*BETAO(J)
      12 SUM = SUM - X(I,J)*BETAO(J)
      11 YNXXB(I) = SUM

C      YNXXB ARE THE RHS OF THE CONSTRAINTS IN THE FORM Y-XB.
C      TEST TO DETERMINE IF PROBLEM P1 OR P2 HAS BEEN
C      SELECTED AND ADJUST YNXXB ACCORDINGLY.
C
      IF (IP1P2.EQ.2) GO TO 15
      DO 16 I=1,NORS
      16 YNXXB(I) = -YNXXB(I)
      15 CONTINUE
      IF(IWRIT1.EQ.0) GO TO 2392
      WRITE(6,60)
      60 FORMAT(1H, ' THE RHS, YNXXB, FOR P1 OR P2')
      DO 61 J=1,NORS
      61 WRITE(6,62) YNXXB(J)
      62 FORMAT(1H, '10X,F14.6)
      2392 CONTINUE

C      CALCULATION OF XBI
C
      DO 13 I=1,IP1
      SUM=0.00
      DO 14 J=1,NORS
      SUM = SUM + (RINV(I,J)+1) - RINV(I,1+NORS+J))*YNXXB(J)
      14 SUM = SUM + (RINV(I,J)+1) - RINV(I,1+NORS+J))*YNXXB(J)
      13 XBI(I) = SUM
      RETURN
      END
      SUBROUTINE XTXINV(X,IP,NORS)
      DIMENSION X(20,10)
      COMMON/XTXIH/XTX(41,41)
      COMMON/IWRIT1,IWRIT2,IWRIT3,IWRIT4,IWRIT5,IWRIT6
      DO 22 I=1,IP
      DO 21 K=1,IP
      SUM=0.00
      DO 20 J=1,NORS
      SUM = SUM + X(J,I)*X(J,K)
      20 CONTINUE
      XTX(I,K)=SUM
      21 CONTINUE
      22 CONTINUE
      CALL INVER,XTX,IP)
      IF (IWRIT1.EQ.0) GO TO 2395
      WRITE(6,67)
      67 FORMAT(1H, ' THE INVERSE OF XTX')
      68 WRITE(6,69) (XTX(I,K),K=1,IP)
      69 FORMAT(1H, '9X,8F14.6)
      2395 CONTINUE
      NOW XTX IS REALLY XTX INVERSE.
      RETURN
      END
      SUBROUTINE RAND(U)
      SUBROUTINE RAND...
      GENERATES UNIFORM (0,1) RANDOM NUMBERS.
      NSEED IS A RANDOM TEN DIGIT INTEGER SUPPLIED IN MAIN.

      IMPLICIT REAL*8 (A-H,O-Z)
      COMMON/NSEED/NSEED
      NP=2147483647
      NL=764261123
      FND=DFLOAT(NP)
      NSEED=NSEED*NHL
      IF (NSEED.LT.0) NSEED=NSEED+NP
      U=DFLOAT(NSEED)
      U=U/FND
      RETURN
      END
      SUBROUTINE NORMAL(NORS,ORS)
      SUBROUTINE NORMAL...
      GENERATES THE EPSILON STAR'S (ERRORS DISTRIBUTED AS
      NORMAL (0,1)) USING A RANDOM NORMAL GENERATOR
      CALLED BUTLER'S ALGORITHM AS ADAPTED FROM RAND.

      DIMENSION RANRNM(20)
      DIMENSION C(6),X(257),U(3),STORE(256),R(256)
      COMMON/NSEED/NSEED
      DATA C/2.515517,.802853,.010328,1.43279,1.89268,.0013081/

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1. REPORT NUMBER 64	2. GOVT ACCESSION NO.	3. RECIPIENT'S CATALOG NUMBER
4. TITLE (and Subtitle) UNBIASED L_1 ESTIMATORS AND THEIR COVARIANCES		5. TYPE OF REPORT & PERIOD COVERED Technical
		6. PERFORMING ORG. REPORT NUMBER
7. AUTHOR(s) Book, D., Booker, J., Hartley, H.O., and Sielken, R.L. Jr.		8. CONTRACT OR GRANT NUMBER(s) N00014-78-C-0426
9. PERFORMING ORGANIZATION NAME AND ADDRESS Texas A&M University Institute of Statistics College Station, Texas 77843		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS NR 047-179
11. CONTROLLING OFFICE NAME AND ADDRESS		12. REPORT DATE June 1980
		13. NUMBER OF PAGES 65
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office) Office of Naval Research		15. SECURITY CLASS. (of this report) Unclassified
		15a. DECLASSIFICATION/DOWNGRADING SCHEDULE
16. DISTRIBUTION STATEMENT (of this Report) Approved for public release; distribution unlimited.		
17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report) NA		
18. SUPPLEMENTARY NOTES		
19. KEY WORDS (Continue on reverse side if necessary and identify by block number) linear regression minimizing the sum of absolute residuals linear programming variances and covariances of L_1 estimators		
20. ABSTRACT (Continue on reverse side if necessary and identify by block number) The parameters in a linear regression model can be estimated by minimizing the sum of the absolute residuals (L_1 estimation) instead of the more classical approach of minimizing the sum of squared residuals (least squares estimation). In addition to other nice properties L_1 estimators are less sensitive to outliers than least squares estimators. This paper describes a linear programming algorithm and computer program		

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S/N 0102-LF-014-6601

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19. Cont.

unbiased L_1 estimators
computer algorithm
least squares
Monte Carlo

20. Cont.

for obtaining unbiased L_1 estimators and estimates of their covariances. These estimated covariances are the new feature in this work and are an extremely important ingredient in hypothesis tests and confidence interval construction. Technical Report 65 provides an analogous treatment of L_1 estimation subject to linear constraints on the parameters.

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